Measuring the Bias of Technological Change*

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Abstract

Technological change can increase the productivity of the various factors of production in equal terms, or it can be biased towards a specific factor. We directly assess the bias of technological change by measuring, at the level of the individual firm, how much of it is labor augmenting and how much is factor neutral. To do so, we develop a framework for estimating production functions when productivity is multi-dimensional. Using panel data from Spain, we find that technological change is biased, with both its labor-augmenting and its factor-neutral components causing output to grow by about 1.5% per year.

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1 Introduction

When technological change occurs, it can increase the productivity of capital, labor, and the other factors of production in equal terms, or it can be biased towards a specific factor. Whether technological change favors some factors of production over others is central to economics. Yet, the empirical evidence is relatively sparse.

The literature on economic growth rests on the assumption that technological change increases the productivity of labor vis-à-vis the other factors of production. It is well known that for a neoclassical growth model to exhibit steady-state growth, either the production function must be Cobb-Douglas or technological change must be labor augmenting (Uzawa 1961), and many endogenous growth models point to human capital accumulation as a source of productivity increases (Lucas 1988, Romer 1990). A number of recent papers provide microfoundations for the literature on economic growth by theoretically establishing that profit-maximizing incentives can ensure that technological change is, at least in the long run, purely labor augmenting (Acemoglu 2003, Jones 2005). Whether this is indeed the case is, however, an empirical question that remains to be answered.

One reason for the scarcity of empirical assessments of the bias of technological change may be a lack of suitable data. Following early work by Brown & de Cani (1963) and David & van de Klundert (1965), economists have estimated aggregate production or cost functions that proxy for labor-augmenting technological change with a time trend (Lucas 1969, Kalt 1978, Antràs 2004, Klump, McAdam & Willman 2007, Binswanger 1974, Cain & Paterson 1986, Jin & Jorgenson 2010).[1] This line of research has produced some evidence of labor-augmenting technological change. However, the intricacies of constructing data series from national income and product accounts (Gordon 1990, Krueger 1999) and the staggering amount of heterogeneity across firms in combination with simultaneously occurring entry and exit (Dunne, Roberts & Samuelson 1988, Davis & Haltiwanger 1992) may make it difficult to interpret a time trend as a meaningful average economy- or sector-wide measure of technological change. Furthermore, this line of research does not provide any deeper insights into the anatomy of the underlying productivity distribution. It also pays scant attention to the fundamental endogeneity problem in production function estimation. This problem arises because a firm's decisions depend on its productivity, and productivity is not observed by the econometrician, and may severely bias the estimates (Marschak & Andrews 1944).[2]

While traditionally using more disaggregated data, the productivity and industrial or-

[1]A much larger literature has estimated the elasticity of substitution using either aggregated or disaggregated data whilst maintaining the assumption of factor-neutral technological change, see Hammermesh (1993) for a survey.

[2]Intuitively, if the firm adjusts to a change in its productivity by expanding or contracting its production, then unobserved productivity and input usage are correlated, resulting in biased estimates of the production function. See Griliches & Mairesse (1998) and Ackerberg, Benkard, Berry & Pakes (2007) for reviews of this and other problems involved in the estimation of production functions.

ganization literatures assume that technological change is factor neutral. Hicks-neutral technological change underlies, either explicitly or implicitly, most of the standard techniques for measuring productivity, ranging from the classic growth decompositions of Solow (1957) and Hall (1988) to the recent structural estimators for production functions that resolve the endogeneity problem (Olley & Pakes 1996, Levinsohn & Petrin 2003, Ackerberg, Caves & Frazer 2015, Doraszelski & Jaumandreu 2013, Gandhi, Navarro & Rivers 2013). In their present form these techniques therefore do not allow us to assess whether technological change is biased towards some factors of production.

In this paper, we combine firm-level panel data that is now widely available with advances in econometric techniques to directly assess the bias of technological change by measuring, at the level of the individual firm, how much of technological change is labor augmenting and how much of it is Hicks neutral. To do so, we develop a framework for estimating production functions when productivity is multi-dimensional and has a labor-augmenting and a Hicks-neutral component.

Our framework accounts for firm-level heterogeneity in the various components of productivity by allowing their evolution to be subject to random shocks. As these productivity innovations accumulate over time, they can cause persistent differences across firms. Because we are able to recover the components of productivity for each firm at each point in time, we obtain a detailed assessment of the impact of technological change at the level it takes place, namely the individual firm. In particular, we are able to assess the dispersion and persistence in the components of productivity and to relate the speed and direction of technological change to firms' R&D activities.

To tackle the endogeneity problem in production function estimation, we build on the insight of Olley & Pakes (1996) that if the decisions that a firm makes can be used to infer its productivity, then productivity can be controlled for in the estimation. We extend their insight to a setting in which productivity is multi- instead of single-dimensional. We infer the firm's productivity from its input usage, in particular its labor and materials decisions. The key insight to identifying the bias of technological change is that Hicks-neutral technological change scales input usage but, in contrast to labor-augmenting technological change, does not change the mix of inputs that a firm uses. A change in the input mix therefore contains information about the bias of technological change, provided we control for the relative prices of the various inputs and other factors that may change the input mix.

We apply the resulting estimator to an unbalanced panel of 2375 Spanish manufacturing firms in ten industries from 1990 to 2006. Spain is an attractive setting for examining the speed and direction of technological change for two reasons. First, Spain became fully integrated into the European Union between the end of the 1980s and the beginning of the 1990s. Any trends in technological change that our analysis uncovers for Spain may thus be viewed as broadly representative for other continental European economies. Second, Spain inherited an industrial structure with few high-tech industries and mostly small and medium-sized firms. R&D is widely seen as lacking (OECD 2007). Yet, Spain grew rapidly during the 1990s, and R&D became increasingly important (European Commission 2001). The accompanying changes in industrial structure are a useful source of variation for analyzing the role of R&D in stimulating different types of technological change.

The particular data set we use has several advantages. The broad coverage allows us to assess the bias of technological change in industries that differ greatly in terms of firms' R&D activities. The data set also has an unusually long time dimension, enabling us to disentangle trends in technological change from short-term fluctuations. Finally, the data set has firm-level prices that we exploit heavily in the estimation.[3]

The Spanish manufacturing sector also poses several challenges for identifying the bias of technological change from a change in the mix of inputs that a firm uses. First, outsourcing directly changes the input mix as the firm procures customized parts and pieces from its suppliers rather than makes them in house from scratch. Second, the Spanish labor market manifestly distinguishes between permanent and temporary labor. We further contribute to the literature following Olley & Pakes (1996) by accounting for outsourcing and the dual nature of the labor market and highlighting the importance of costly adjustments to permanent labor for measuring the bias of technological change.

Our estimates provide clear evidence that technological change is biased. Ceteris paribus labor-augmenting technological change causes output to grow, on average, in the vicinity of 1.5% per year. While there is a shift from unskilled to skilled workers in our data, this skill upgrading explains some but not all of the growth of labor-augmenting productivity. In many industries, labor-augmenting productivity grows because workers with a given set of skills become more productive over time.

At the same time, our estimates show that Hicks-neutral technological change plays an equally important role. In addition to labor-augmenting technological change, Hicks-neutral technological change causes output to grow, on average, in the vicinity of 1.5% per year.

Behind these averages lies a substantial amount of heterogeneity across industries and firms. Our estimates point to substantial and persistent differences in labor-augmenting and Hicks-neutral productivity across firms, in line with the "stylized facts" about productivity in Bartelsman & Doms (2000) and Syverson (2011). Beyond these facts, we show that, at the level of the individual firm, the levels of labor-augmenting and Hicks-neutral productivity are positively correlated, as are their rates of growth.

Our estimates further indicate that firms' R&D activities play a key role in determining the differences in the components of productivity across firms and their evolution over time. Interestingly, labor-augmenting productivity is slightly more closely tied to firms' R&D activities than is Hicks-neutral productivity. Through the lens of the literature on

[3] There are other firm-level data sets such as the Colombian Annual Manufacturers Survey (Eslava, Haltiwanger, Kugler & Kugler 2004) and the Longitudinal Business Database at the U.S. Census Bureau that contain separate information on prices and quantities, at least for a subset of industries (Roberts & Supina 1996, Foster, Haltiwanger & Syverson 2008, Foster, Haltiwanger & Syverson 2016).

induced innovation and directed technical change (Hicks 1932, Acemoglu 2002), this may be viewed as supporting the argument that firms direct their R&D activities to conserve on labor.

Biased technological change has consequence far beyond the growth of output. To illustrate, we use our estimates to show that biased technological change is the primary driver of the decline of the aggregate share of labor in the Spanish manufacturing sector over our sample period. Similar declines have been observed in many advanced economies in past decades and have attracted considerable attention in the macroeconomics literature (Blanchard 1997, Bentolila & Saint-Paul 2004, McAdam & Willman 2013, Karabarbounis & Neiman 2014, Oberfield & Raval 2014).

The starting point of this paper is the literature on the structural estimation of production functions. Olley & Pakes (1996), Levinsohn & Petrin (2003), Ackerberg et al. (2015), and many others specify a Cobb-Douglas production function. Productivity is single-dimensional or, equivalently, technological change is Hicks neutral by construction.[4] To assess the bias of technological change, we generalize the Cobb-Douglas production function and allow productivity to be multi-dimensional.

Our approach further differs from much of the previous literature by exploiting the parameter restrictions between the production and input demand functions, as in Doraszelski & Jaumandreu (2013). This allows us to parametrically invert from observed input usage to unobserved productivity and eases the demands on the data compared to the nonparametric inversion in Olley & Pakes (1996), Levinsohn & Petrin (2003), and Ackerberg et al. (2015), especially if the input demand functions are high-dimensional.[5]

Our paper is related to Van Biesebroeck (2003). Using plant-level panel data for the U.S. automobile industry, he estimates a plant's Hicks-neutral productivity as a fixed effect and parametrically inverts from its input usage to the plant's capital-biased (also called labor-saving) productivity. Our approach is more general in that we allow all components of productivity to evolve over time and in response to firms' R&D activities.

Our paper is also related to Grieco, Li & Zhang (2016). Because their data contains the materials bill rather than its split into price and quantity, they parametrically invert from its input usage to a firm's Hicks-neutral productivity and the price of materials that the firm faces. In subsequent work in progress, Zhang (2014, 2015) applies the same idea to infer a firm's capital- and labor-augmenting productivity.[6]

Finally, our paper touches—although more tangentially—on the literature on skill bias

[4]As is well known, a Cobb-Douglas production function has an elasticity of substitution of one and therefore cannot be used to separate different types of technological change. Our data rejects a Cobb-Douglas production function (see Section 6).

[5]See Doraszelski & Jaumandreu (2013) for details on the pros and cons of the parametric inversion.

[6]While Grieco et al. (2016) and Zhang (2014, 2015) build on Doraszelski & Jaumandreu (2013) by exploiting the parameter restrictions between the production and input demand functions, they differ by plugging the recovered unobservables back into the production function. This avoids assumptions on the law of motion for productivity. However, parameters of interest may cancel depending on the specification of the production function (see Section 2.1 of Grieco et al. (2016) and Section 3 of Ackerberg et al. (2015)).

that studies the differential impact of technological change, especially in the form of computerization, on the various types of labor. Our approach is similar to some of the recent work on skill bias (Machin & Van Reenen 1998, Black & Lynch 2001, Abowd, Haltiwanger, Lane, McKinney & Sandusky 2007, Bloom, Sadun & Van Reenen 2012) in that it starts from a production function and focuses on the individual firm. While we focus on labor versus the other factors of production, the techniques we develop may be adapted to investigate the skill bias of technological change, although our particular data set is not ideal for this purpose.

The remainder of this paper is organized as follows: Section 2 explains how we identify the bias of technological change and previews our empirical strategy. Section 3 describes the data and some patterns in the data that inform the subsequent analysis. Section 4 sets out a dynamic model of the firm. Section 5 develops an estimator for production functions when productivity is multi-dimensional. Sections 6–9 present our main results on labor-augmenting and Hicks-neutral technological change. Section 10 explores whether capital-augmenting technological change plays a role in our data in addition to labor-augmenting and Hicks-neutral technological change. Section 11 concludes.

Throughout the paper, we adopt the convention that uppercase letters denote levels and lowercase letters denote logs. Unless noted otherwise, we refer to output and the various factors of production in terms of quantity and not in terms of value. In particular, we refer to the value of labor as the wage bill and to the value of materials as the materials bill.

2 Labor-augmenting and Hicks-neutral productivity

We infer a firm's productivity from its input usage, in particular its labor and materials decisions. To separate labor-augmenting from Hicks-neutral productivity, we exploit that the input mix is closely related to—and therefore contains information about—labor-augmenting productivity. We also show that the constant elasticity of substitution (CES) production function that we use in our application approximates, to a first order, the relationship between the input mix and labor-augmenting productivity that arises in a wider class of production functions. To facilitate the exposition, we proceed in a highly simplified setting. Our application extends the setting to accommodate the institutional realities of the Spanish manufacturing sector.

Consider a firm with the production function

$$Y_{jt} = F(K_{jt}, \exp(\omega_{Ljt})L_{jt}, M_{jt}) \exp(\omega_{Hjt}) \exp(e_{jt}), \tag{1}$$

where Y_{jt} is the output of firm j in period t, K_{jt} is capital, L_{jt} is labor, and M_{jt} is materials. The labor-augmenting productivity of firm j in period t is ω_{Ljt} and its Hicks-neutral productivity is ω_{Hjt}. The production function in equation (1) abstracts from capital-augmenting productivity for reasons explained in Sections 3 and 10. Finally, e_{jt} is a random shock.

To relate the input ratio $\frac{M_{jt}}{L_{jt}}$ to labor-augmenting productivity ω_{Ljt}, we assume that $(\exp(\omega_{Ljt})L_{jt}, M_{jt})$ is separable from K_{jt} in that the function $F(\cdot)$ in equation (1) is composed of the functions $G(\cdot)$ and $H(\cdot)$ as

$$F(K_{jt}, \exp(\omega_{Ljt})L_{jt}, M_{jt}) = G(K_{jt}, H(\exp(\omega_{Ljt})L_{jt}, M_{jt})), \tag{2}$$

where $H(\exp(\omega_{Ljt})L_{jt}, M_{jt})$ is homogeneous of arbitrary degree.[7] Without loss of generality, we set the degree of homogeneity to one. Throughout we maintain that all functions are differentiable as needed. As in Levinsohn & Petrin (2003), we finally assume that labor and materials are static (or "variable") inputs that are chosen each period to maximize short-run profits and that the firm is a price-taker in input markets, where it faces W_{jt} and P_{Mjt} as prices of labor and materials, respectively.

The input ratio $\frac{M_{jt}}{L_{jt}}$ is therefore the solution to the ratio of the first-order conditions for labor and materials

$$\frac{\frac{\partial H\left(\exp(\omega_{Ljt})L_{jt}, M_{jt}\right)}{\partial L_{jt}} \exp(\omega_{Ljt})}{\frac{\partial H\left(\exp(\omega_{Ljt})L_{jt}, M_{jt}\right)}{\partial M_{jt}}} = \frac{\frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial L_{jt}} \exp(\omega_{Ljt})}{\frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial M_{jt}}} = \frac{W_{jt}}{P_{Mjt}}, \tag{3}$$

where the first equality uses that $H(\exp(\omega_{Ljt})L_{jt}, M_{jt})$ is homogeneous of degree one and, recall, uppercase letters denote levels and lowercase letters denote logs.

Equation (3) implies that the input ratio $\frac{M_{jt}}{L_{jt}}$ depends on the price ratio $\frac{P_{Mjt}}{W_{jt}}$ and labor-augmenting productivity ω_{Ljt}. Importantly, the input ratio $\frac{M_{jt}}{L_{jt}}$ does not depend on Hicks-neutral productivity ω_{Hjt}. This formalizes that the mix of inputs that a firm uses is related to—and therefore contains information about—its labor-augmenting productivity but is unrelated to its Hicks-neutral productivity. Intuitively, the labor and materials decisions hinge on the marginal products of labor and materials. Because the marginal products are proportional to Hicks-neutral productivity, materials per unit of labor as determined by the ratio of the first-order conditions in equation (3) is unrelated to Hicks-neutral productivity. In this sense, separating labor-augmenting from Hicks-neutral productivity does not rely on functional form beyond the separability assumption in equation (2).[8]

The following proposition further characterizes the log of the input ratio $m_{jt} - l_{jt}$:

Proposition 1 *The input ratio $m_{jt} - l_{jt}$ has the first-order Taylor series*

$$\gamma_L^0 - \sigma\left(\exp(\omega_{Ljt}^0 - (m_{jt}^0 - l_{jt}^0))\right)(p_{Mjt} - w_{jt}) + \left(1 - \sigma\left(\exp(\omega_{Ljt}^0 - (m_{jt}^0 - l_{jt}^0))\right)\right)\omega_{Ljt} \tag{4}$$

[7] Equation (2) immediately implies that $F(K_{jt}, \exp(\omega_{Ljt})L_{jt}, M_{jt})$ is weakly separable in the partition $(K_{jt}, (\exp(\omega_{Ljt})L_{jt}, M_{jt}))$ (Chambers 1988, equation (1.26)). It is equivalent to $F(K_{jt}, \exp(\omega_{Ljt})L_{jt}, M_{jt})$ being weakly separable under some additional monotonicity and quasi-concavity assumptions (Goldman & Uzawa 1964).

[8] One can forgo the separability assumption by relying more on functional form. Our empirical strategy generalizes to, for example, a translog production function that does not satisfy equation (2).

around a point $(m^0_{jt} - l^0_{jt}, p^0_{Mjt} - w^0_{jt}, \omega^0_{Ljt})$ *satisfying equation (3), where* γ^0_L *is a constant and* $\sigma\left(\exp(\omega^0_{Ljt} - (m^0_{jt} - l^0_{jt}))\right)$ *is the elasticity of substitution between materials and labor in the production function in equation (1).*

The proof can be found in Appendix A.

Our application uses a CES production function

$$Y_{jt} = \left[\beta_K K_{jt}^{-\frac{1-\sigma}{\sigma}} + (\exp(\omega_{Ljt})L_{jt})^{-\frac{1-\sigma}{\sigma}} + \beta_M M_{jt}^{-\frac{1-\sigma}{\sigma}}\right]^{-\frac{\nu\sigma}{1-\sigma}} \exp(\omega_{Hjt})\exp(e_{jt}),$$

where ν and σ are the elasticity of scale and substitution, respectively, and β_K and β_M are the so-called distributional parameters.[9] Depending on the elasticity of substitution, the CES production function encompasses the special cases of a Leontieff ($\sigma \to 0$), Cobb-Douglas ($\sigma = 1$), and linear ($\sigma \to \infty$) production function.

The ratio of the first-order conditions in equation (3) implies

$$m_{jt} - l_{jt} = \sigma \ln \beta_M - \sigma(p_{Mjt} - w_{jt}) + (1-\sigma)\omega_{Ljt}. \tag{5}$$

Comparing equations (4) and (5) shows that the CES production function approximates, to a first order, the input ratio $m_{jt} - l_{jt}$ arising from an arbitrary production function satisfying equation (2). This gives a sense of robustness to the CES production function.[10]

Our empirical strategy uses equation (5) to infer a firm's labor-augmenting productivity from its input mix. In doing so, we must control for other factors besides the relative prices of the various inputs that may change the input mix, in particular outsourcing and the dual nature of the Spanish labor market. With labor-augmenting productivity in hand, we use the first-order condition for labor to recover Hicks-neutral productivity. The remainder of our empirical strategy follows along the lines of Olley & Pakes (1996), Levinsohn & Petrin (2003), Ackerberg et al. (2015), and Doraszelski & Jaumandreu (2013) by combining the inferred productivities with their laws of motion to set up estimation equations.

Equation (5) has a long tradition in the literature, although it is used in a very different way from ours. With skilled and unskilled workers in place of materials and labor, equation (5) is at the heart of the literature on skill bias (see Card & DiNardo (2002) and Violante (2008) and the references therein). With capital in place of materials, equation (5) serves

[9] We implicitly set the constant of proportionality β_0 to one because it cannot be separated from an additive constant in Hicks-neutral productivity ω_{Hjt}. We estimate them jointly and carefully ensure that the reported results depend only on their sum. We similarly normalize the distributional parameter β_L. Technological change can therefore equivalently be thought of as letting these parameters of the production function vary by firm and time. The nascent literature on heterogeneous production functions (Balat, Brambilla & Sasaki 2015, Fox, Hadad, Hoderlein, Petrin & Sherman 2016, Kasahara, Schrimpf & Suzuki 2015) explores to what extent it is possible to let all parameters of the production function vary by firm and time.

[10] It also suggests that our "nonparametric" estimates of labor-augmenting technological change can be fed into a growth decomposition along the lines of Solow (1957) and Hall (1988) to obtain a "nonparametric" estimate of Hicks-neutral technological change. We leave this to future research.

to estimate the elasticity of substitution σ in an aggregate value-added production function (see Antràs 2004). More recently, Raval (2013) uses a variant of equation (5) obtained from a value-added production function with capital- and labor-augmenting productivity to estimate σ from firm-level panel data.[11]

Equation (5) is typically estimated by OLS. The problem is that labor-augmenting productivity, which is not observed by the econometrician, is correlated over time and also with the wage. We intuitively expect the wage to be higher when labor is more productive, even if it adjusts slowly with some lag. This positive correlation induces an upward bias in the estimate of the elasticity of substitution. This is a variant of the endogeneity problem in production function estimation. Because we use equation (5) to recover labor-augmenting productivity rather than directly estimate it, we are able tackle the endogeneity problem with a combination of assumptions on the timing of decisions and the evolution of the components of productivity.

3 Data

Our data comes from the Encuesta Sobre Estrategias Empresariales (ESEE) survey, a firm-level survey of the Spanish manufacturing sector sponsored by the Ministry of Industry. The unit of observation is the firm, not the plant or the establishment. Our data covers the 1990s and early 2000s. At the beginning of the survey in 1990, 5% of firms with up to 200 workers were sampled randomly by industry and size strata. All firms with more than 200 workers were asked to participate in the survey, and 70% of them complied. Some firms vanish from the sample due to either exit (shutdown by death or abandonment of activity) or attrition. These reasons can be distinguished in the data and attrition remained within acceptable limits. To preserve representativeness, newly created firms were added to the sample every year. We provide details on industry and variable definitions in Appendix B.

Our sample covers a total of 2375 firms in ten industries when restricted to firms with at least three years of data. Columns (1) and (2) of Table 1 show the number of observations and firms by industry. Sample sizes are moderate. Newly created firms are a large fraction of the total number of firms, ranging from 26% to 50% in the different industries. There is a much smaller fraction of exiting firms, ranging from 6% to 15% and above in a few industries. Firms remain in the sample from a minimum of three years to a maximum of 16 years between 1990 and 2006.

The 1990s and early 2000s were a period of rapid output growth, coupled with stagnant or, at best, slightly increasing employment and intense investment in physical capital, see columns (3)–(6) of Table 1. Consistent with this rapid growth, firms on average report that their markets are slightly more often expanding rather than contracting; hence, demand

[11] These latter forms of equation (5) rest on the assumption that capital is a static input that is chosen each period to maximize short-run profits. In contrast, the literature following Olley & Pakes (1996) stresses that the choice of capital has dynamic implications.

tends to shift out over time.

An attractive feature of our data is that it contains firm-specific, Paasche-type price indices for output and materials. We note that the variation in these price indices is partly due to changes over time in the bundles of goods that make up output and, respectively, materials (see Bernard, Redding & Schott (2010) and Goldberg, Khandelwal, Pavcnik & Topalova (2010) for evidence on product turnover), and that these changes may be related to a firm's productivity. The growth of prices, averaged from the growth of prices as reported individually by each firm, is moderate. The growth of the price of output in column (7) ranges from 0.8% to 2.1%. The growth of the wage ranges from 4.3% to 5.4% and the growth of the price of materials ranges from 2.8% to 4.1%.

Biased technological change. The evolution of the relative quantities and prices of the various factors of production already hints at an important role for labor-augmenting technological change. As columns (8) and (9) of Table 1 show, with the exception of industries 7, 8, and 9, the input ratio $\frac{M_{jt}}{L_{jt}}$ increases much more than the price ratio $\frac{P_{Mjt}}{W_{jt}}$ decreases. One possible explanation is that the elasticity of substitution between materials and labor exceeds 1. To see this, recall that the elasticity of substitution (Chambers 1988, equation (1.12)) is

$$\frac{d\ln\left(\frac{M_{jt}}{L_{jt}}\right)}{d\ln\left(|MRTS_{MLjt}|\right)} = -\frac{d\ln\left(\frac{M_{jt}}{L_{jt}}\right)}{d\ln\left(\frac{P_{Mjt}}{W_{jt}}\right)},$$

where $|MRTS_{MLjt}|$ is the absolute value of the marginal rate of technological substitution between materials and labor, and the equality follows to the extent that it equals the price ratio $\frac{P_{Mjt}}{W_{jt}}$. However, because the estimates of the elasticity of substitution in the previous literature lie somewhere between 0 and 1 (see Chirinko (2008) and the references therein for the elasticity of substitution between capital and labor and Bruno (1984), Rotemberg & Woodford (1996), and Oberfield & Raval (2014) for the elasticity of substitution between materials and an aggregate of capital and labor), this explanation is implausible. Labor-augmenting technological change offers an alternative explanation. As it makes labor more productive, equation (4) implies that it directly increases materials per unit of labor. Thus, labor-augmenting technological change may go a long way in rationalizing why the change in the input ratio $\frac{M_{jt}}{L_{jt}}$ exceeds the change in the price ratio $\frac{P_{Mjt}}{W_{jt}}$.

In contrast, columns (10) and (11) of Table 1 provide no evidence for capital-augmenting technological change. The investment boom in Spain in the 1990s and early 2000s was fueled by improved access to European and international capital markets. With the exception of industries 5, 6, and 8, the concomitant decrease in the input ratio $\frac{M_{jt}}{K_{jt}}$ is much smaller than the increase in the price ratio $\frac{P_{Mjt}}{P_{Kjt}}$, where P_{Kjt} is the price of capital as measured by the user cost in our data.[12] This pattern is consistent with an elasticity of substitution between

[12] The user cost is a notably rough measure of the price of capital. In particular, the price of capital includes adjustment costs, and as a shadow price, it is unobservable. The user cost, in contrast, is based

materials and capital between 0 and 1. Indeed, capital-augmenting technological change can only directly contribute to the decline in materials per unit of capital in the unlikely scenario that it makes capital less productive.

Based on these patterns in the data we focus on labor-augmenting technological change in the subsequent analysis. We return to capital-augmenting technological change in Section 10. In the remainder of this section we point out other features of the data that figure prominently in our analysis.

Temporary labor. In recognition of the dual nature of the Spanish labor market, we distinguish between permanent and temporary labor. We treat temporary labor as a static input that is chosen each period to maximize short-run profits. This is appropriate because Spain greatly enhanced the possibilities for hiring and firing temporary workers during the 1980s, and by the beginning of the 1990s had one the highest shares of temporary workers in Europe (Dolado, Garcia-Serrano & Jimeno 2002). Temporary workers are employed for fixed terms with no or very small severance pay. In our sample, between 72% and 84% of firms use temporary labor and, among the firms that do, its share of the labor force ranges from 16% in industry 10 to 32% in industry 9, see columns (1) and (2) of Table 2.

Rapid expansions and contractions of temporary labor are common: The difference between the maximum and the minimum share of temporary labor within a firm ranges on average from 20% to 33% across industries (column (3)). In addition to distinguishing temporary from permanent labor, we measure labor as hours worked (see Appendix B). At this margin, firms enjoy a high degree of flexibility: Within a firm, the difference between the maximum and the minimum hours worked ranges on average from 43% to 56% across industries, and the difference between the maximum and the minimum hours per worker ranges on average from 4% to 13% (columns (4) and (5)).

Outsourcing. We account for outsourcing in our analysis. Outsourcing may directly contribute to the shift from labor to materials that column (8) of Table 1 documents as firms procure customized parts and pieces from their suppliers rather than make them in house from scratch. As can be seen in columns (6) and (7) of Table 2, between 21% and 57% of firms in our sample engage in outsourcing. Among the firms that do, the share of outsourcing in the materials bill ranges from 14% in industry 7 to 29% in industry 4. While the share of outsourcing remains stable over our sample period, the standard deviation in column (7) indicates a substantial amount of heterogeneity across the firms within an industry, similar to the share of temporary labor in column (2).

Firms' R&D activities. The R&D intensity of Spanish manufacturing firms is low by European standards, but R&D became increasingly important during the 1990s (see, e.g.,

solely on observables (see Appendix B).

European Commission 2001).[13] Columns (8)–(10) of Table 2 show that the ten industries differ markedly in terms of firms' R&D activities and that there is again substantial heterogeneity across the firms within an industry. Industries 3, 4, 5, and 6 exhibit high innovative activity. More than two-thirds of firms perform R&D during at least one year in the sample period, with at least 36% of stable performers engaging in R&D in all years (column (8)) and at least 28% of occasional performers engaging in R&D in some but not all years (column (9)). The R&D intensity among performers ranges on average from 2.2% to 2.9% (column (10)). Industries 1, 2, 7, and 8 are in an intermediate position. Less than half of firms perform R&D, and there are fewer stable than occasional performers. The R&D intensity is on average between 1.1% and 1.7% with a much lower value of 0.7% in industry 7. Finally, industries 9 and 10 exhibit low innovative activity. About a third of firms perform R&D, and the R&D intensity is on average between 1.0% and 1.5%.

4 A dynamic model of the firm

The purpose of our model is to enable us to infer a firm's productivity from its input usage and to clarify our assumptions on the timing of decisions that we rely on in estimation. We extend the previous literature on the structural estimation of production functions by allowing productivity to be multi-dimensional. We further contribute to the literature following Olley & Pakes (1996) by accounting for outsourcing and the dual nature of the Spanish labor market.

Production function. The firm has the CES production function

$$Y_{jt} = \left[\beta_K K_{jt}^{-\frac{1-\sigma}{\sigma}} + \left(\exp(\omega_{Ljt}) L_{jt}^*\right)^{-\frac{1-\sigma}{\sigma}} + \beta_M \left(M_{jt}^*\right)^{-\frac{1-\sigma}{\sigma}}\right]^{-\frac{\nu\sigma}{1-\sigma}} \exp(\omega_{Hjt}) \exp(e_{jt}), \quad (6)$$

where Y_{jt} is the output of firm j in period t, K_{jt} is capital, ω_{Ljt} is labor-augmenting productivity, ω_{Hjt} is Hicks-neutral productivity, and e_{jt} is a mean zero random shock that is uncorrelated over time and across firms. Extending the setting in Section 2, $L_{jt}^* = \Lambda(L_{Pjt}, L_{Tjt})$ is an aggregate of permanent labor L_{Pjt} and temporary labor L_{Tjt} and $M_{jt}^* = \Gamma(M_{Ijt}, M_{Ojt})$ is an aggregate of in-house materials M_{Ijt} and outsourced materials (customized parts and pieces) M_{Ojt}. The aggregators $\Lambda(L_{Pjt}, L_{Tjt})$ and $\Gamma(M_{Ijt}, M_{Ojt})$ accommodate differences in the productivities of permanent and temporary labor, respectively, in-house and outsourced materials; we do not further specify these aggregators.

The production function in equation (6) is the most parsimonious we can use to separate labor-augmenting from Hicks-neutral productivity. It encompasses three restrictions. First, technological change does not affect the parameters ν and σ, as we are unaware of evidence suggesting that the elasticity of scale or the elasticity of substitution varies over our sample

[13] R&D intensities for manufacturing firms are 2.1% in France, 2.6% in Germany, and 2.2% in the UK as compared to 0.6% in Spain (European Commission 2004).

period. Second, the elasticity of substitution between capital, labor, and materials is the same.[14] We assess this restriction in Section 8. For now we note that previous estimates of the elasticity of substitution between materials and an aggregate of capital and labor (Bruno 1984, Rotemberg & Woodford 1996, Oberfield & Raval 2014) fall in the same range as estimates of the elasticity of substitution between capital and labor (Chirinko 2008). Third, the productivities of capital and materials are restricted to change at the same rate and in lockstep with Hicks-neutral technological change.[15] Treating capital and materials the same is in line with the fact that both are, at least to a large extent, produced goods. In contrast, labor is traditionally viewed as unique among the various factors of production (Marshall 1920), and changes in its productivity are a tenet of the literature on economic growth. The patterns in the data described in Section 3 further justify focusing on labor-augmenting technological change. In Section 10, we explore more thoroughly whether capital-augmenting technological change plays a role in our data in addition to labor-augmenting and Hicks-neutral technological change.

Laws of motion. The components of productivity are presumably correlated with each other and over time and possibly also correlated across firms. As in Doraszelski & Jaumandreu (2013), we endogenize productivity by incorporating R&D expenditures into the model. We assume that the evolution of the components of productivity is governed by controlled first-order, time-inhomogeneous Markov processes with transition probabilities $P_{Lt+1}(\omega_{Ljt+1}|\omega_{Ljt}, R_{jt})$ and $P_{Ht+1}(\omega_{Hjt+1}|\omega_{Hjt}, R_{jt})$, where R_{jt} is R&D expenditures. Despite their parsimony, these stochastic processes accommodate correlation between the components of productivity.[16] Moreover, because they are time-inhomogeneous, they accommodate secular trends in productivity.

The firm knows its current productivity when it makes its decisions for period t and anticipates the effect of R&D on its future productivity. The Markovian assumption implies

$$\omega_{Ljt+1} = E_t\left[\omega_{Ljt+1}|\omega_{Ljt}, R_{jt}\right] + \xi_{Ljt+1} = g_{Lt}(\omega_{Ljt}, R_{jt}) + \xi_{Ljt+1}, \tag{7}$$

$$\omega_{Hjt+1} = E_t\left[\omega_{Hjt+1}|\omega_{Hjt}, R_{jt}\right] + \xi_{Hjt+1} = g_{Ht}(\omega_{Hjt}, R_{jt}) + \xi_{Hjt+1}. \tag{8}$$

That is, *actual* labor-augmenting productivity ω_{Ljt+1} in period $t+1$ decomposes into *ex-*

[14] The elasticity of substitution between L_{Pjt} and L_{Tjt}, respectively, M_{Ijt} and M_{Ojt} depends on the aggregators $\Lambda(L_{Pjt}, L_{Tjt})$ and $\Gamma(M_{Ijt}, M_{Ojt})$ and may differ from σ.

[15] A production function with capital-augmenting, labor-augmenting, and materials-augmenting productivity that is homogeneous of arbitrary degree is equivalent to a production function with capital-augmenting, labor-augmenting, and Hicks-neutral productivity. Without loss of generality, we therefore subsume the common component of capital-augmenting, labor-augmenting, and materials-augmenting technological change into Hicks-neutral productivity.

[16] Our empirical strategy generalizes to a joint Markov process $P_{t+1}(\omega_{Ljt+1}, \omega_{Hjt+1}|\omega_{Ljt}, \omega_{Hjt}, r_{jt})$. While R&D is widely seen as a major source of productivity growth (Griliches 1998, Griliches 2000), our empirical strategy extends to other sources such as technology adoption, learning-by-importing (Kasahara & Rodrigue 2008), and learning-by-exporting (De Loecker 2013). Both extensions are demanding on the data, however, as they increase the dimensionality of the functions that must be nonparametrically estimated.

pected labor-augmenting productivity $g_{Lt}(\omega_{Ljt}, R_{jt})$ and a random shock ξ_{Ljt+1}. This productivity innovation is by construction mean independent (although not necessarily fully independent) of ω_{Ljt} and R_{jt}. It captures the uncertainties that are naturally linked to productivity as well as those that are inherent in the R&D process such as chance of discovery, degree of applicability, and success in implementation. Nonlinearities in the link between R&D and productivity are captured by the conditional expectation function $g_{Lt}(\cdot)$ that we estimate nonparametrically along with the parameters of the production function. Actual Hicks-neutral productivity ω_{Hjt+1} decomposes similarly.

Capital accumulates according to $K_{jt+1} = (1 - \delta)K_{jt} + I_{jt}$, where δ is the rate of depreciation. As in Olley & Pakes (1996), investment I_{jt} chosen in period t becomes effective in period $t + 1$. Choosing I_{jt} is therefore equivalent to choosing K_{jt+1}.

In recognition of the dual nature of the Spanish labor market, we distinguish between permanent and temporary labor. Permanent labor is subject to convex adjustment costs $C_{L_P}(L_{Pjt}, L_{Pjt-1})$ that reflect the substantial cost of hiring and firing that the firm may incur (Hammermesh 1993, Hammermesh & Pfann 1996). The choice of permanent labor thus may have dynamic implications. In contrast, temporary labor is a static input.

We further distinguish between in-house and outsourced materials. Outsourcing is, to a large extent, based on contractual relationships between the firm and its suppliers (Grossman & Helpman 2002, Grossman & Helpman 2005). The ratio of outsourced to in-house materials $Q_{Mjt} = \frac{M_{Ojt}}{M_{Ijt}}$ is subject to (convex or not) adjustment costs $C_{Q_M}(Q_{Mjt+1}, Q_{Mjt})$ that stem from forming and dissolving these relationships. The firm must maintain Q_{Mjt} but may scale M_{Ijt} and M_{Ojt} up or down at will; in-house materials, in particular, is a static input. In the Online Appendix, we develop an alternative model of outsourcing that assumes that both in-house and outsourced materials are static inputs that the firm may mix and match at will, thereby dispensing with the costly-to-adjust ratio of outsourced to in-house materials.

Output and input markets. The firm has market power in the output market, e.g., because products are differentiated. Its inverse residual demand function $P(Y_{jt}, D_{jt})$ depends on its output Y_{jt} and the demand shifter D_{jt}.[17] The firm is a price-taker in input markets, where it faces W_{Pjt}, W_{Tjt}, P_{Ijt}, and P_{Ojt} as prices of permanent and temporary labor and in-house and outsourced materials, respectively. In Section 6 we instead assume that the firm faces a menu of qualities and wages in the market for permanent labor.

The demand shifter and the prices that the firm faces in input markets evolve according to a Markov process that we do not further specify. As a consequence, the prices that the firm faces in period t+1 may depend on its productivity in period t or on an average industry-wide measure of productivity. Finally, the Markov process may be time-inhomogenous to

[17] In general, the residual demand that the firm faces depends on its rivals' prices. In taking the model to the data, one may replace rivals' prices by an aggregate price index or dummies, although this substantially increases the dimensionality of the functions that must be nonparametrically estimated.

accommodate secular trends.

Bellman equation. The firm makes its decisions in a discrete-time setting with the goal of maximizing the expected net present value of future cash flows. In contrast to its labor-augmenting productivity ω_{Ljt} and its Hicks-neutral productivity ω_{Hjt}, the firm does not know the random shock e_{jt} when it makes its decisions for period t. Letting $V_t(\cdot)$ denote the value function in period t, the Bellman equation for the firm's dynamic programming problem is

$$\begin{aligned} V_t(\Omega_{jt}) \quad = \quad & \max_{K_{jt+1}, L_{Pjt}, L_{Tjt}, Q_{Mjt+1}, M_{Ijt}, R_{jt}} P\left(X_{jt}^{-\frac{\nu\sigma}{1-\sigma}} \exp(\omega_{Hjt}), D_{jt}\right) X_{jt}^{-\frac{\nu\sigma}{1-\sigma}} \exp(\omega_{Hjt})\mu \\ & - C_I(K_{jt+1} - (1-\delta)K_{jt}) - W_{Pjt}L_{Pjt} - C_{L_P}(L_{Pjt}, L_{Pjt-1}) - W_{Tjt}L_{Tjt} \\ & - (P_{Ijt} + P_{Ojt}Q_{Mjt}) M_{Ijt} - C_{Q_M}(Q_{Mjt+1}, Q_{Mjt}) - C_R(R_{jt}) \\ & + \frac{1}{1+\rho} E_t \left[V_{t+1}(\Omega_{jt+1}) | \Omega_{jt}, R_{jt}\right], \end{aligned} \tag{9}$$

where

$$X_{jt} = \beta_K K_{jt}^{-\frac{1-\sigma}{\sigma}} + \left(\exp(\omega_{Ljt}) L_{jt}^*\right)^{-\frac{1-\sigma}{\sigma}} + \beta_M \left(M_{jt}^*\right)^{-\frac{1-\sigma}{\sigma}}, \quad \mu = E_t\left[\exp(e_{jt})\right],$$

$\Omega_{jt} = (K_{jt}, L_{Pjt-1}, Q_{Mjt}, \omega_{Ljt}, \omega_{Hjt}, W_{Pjt}, W_{Tjt}, P_{Ijt}, P_{Ojt}, D_{jt})$ is the vector of state variables, and ρ is the discount rate. $C_I(I_{jt})$ and $C_R(R_{jt})$ are the cost of investment and R&D, respectively, and accommodate indivisibilities in investment and R&D projects. The firm's dynamic programming problem gives rise to policy functions that characterize its investment and R&D decisions (and thus the values of K_{jt+1} or, equivalently, I_{jt} and R_{jt} in period t) as well as its input usage (L_{Pjt}, L_{Tjt}, Q_{Mjt+1}, and M_{Ijt}). The latter is central to our empirical strategy.

Investment and R&D decisions. The investment and R&D decisions depend on the vector of state variables in our model. In the spirit of the literature on induced innovation, the firm may account for current input prices (as they are part of Ω_{jt}) and its expectation of future input prices (through the continuation value in equation (9)).[18]

[18] The firm may further account for its expectation of future output demand and input supply conditions. Because our empirical strategy infers the firm's productivity from its labor and materials decisions, it is not affected by including additional state variables to model the evolution of these conditions in our model besides the demand shifter D_{jt}.

Input usage. We infer the firm's productivity from its labor and materials decisions. The first-order conditions for permanent and temporary labor are

$$\nu\mu X_{jt}^{-\left(1+\frac{\nu\sigma}{1-\sigma}\right)} \exp\left(\omega_{Hjt}\right) \exp\left(-\frac{1-\sigma}{\sigma}\omega_{Ljt}\right) \left(L_{jt}^{*}\right)^{-\frac{1}{\sigma}} \frac{\partial L_{jt}^{*}}{\partial L_{Pjt}} = \frac{W_{Pjt}(1+\Delta_{jt})}{P_{jt}\left(1-\frac{1}{\eta(p_{jt},D_{jt})}\right)}, \quad (10)$$

$$\nu\mu X_{jt}^{-\left(1+\frac{\nu\sigma}{1-\sigma}\right)} \exp\left(\omega_{Hjt}\right) \exp\left(-\frac{1-\sigma}{\sigma}\omega_{Ljt}\right) \left(L_{jt}^{*}\right)^{-\frac{1}{\sigma}} \frac{\partial L_{jt}^{*}}{\partial L_{Tjt}} = \frac{W_{Tjt}}{P_{jt}\left(1-\frac{1}{\eta(p_{jt},D_{jt})}\right)}, \quad (11)$$

where $\eta(p_{jt}, D_{jt})$ is the absolute value of the price elasticity of the residual demand that the firm faces, and by the envelope theorem, the gap between the wage of permanent workers W_{Pjt} and the shadow wage is

$$\begin{aligned}
\Delta_{jt} &= \frac{1}{W_{Pjt}}\left(\frac{\partial C_{L_P}(L_{Pjt}, L_{Pjt-1})}{\partial L_{Pjt}} - \frac{1}{1+\rho}E_t\left[\frac{\partial V_{t+1}(\Omega_{jt+1})}{\partial L_{Pjt}}|\Omega_{jt}, R_{jt}\right]\right) \\
&= \frac{1}{W_{Pjt}}\left(\frac{\partial C_{L_P}(L_{Pjt}, L_{Pjt-1})}{\partial L_{Pjt}} + \frac{1}{1+\rho}E_t\left[\frac{\partial C_{L_P}(L_{Pjt+1}, L_{Pjt})}{\partial L_{Pjt}}|\Omega_{jt}, R_{jt}\right]\right).
\end{aligned}$$

Equations (10) and (11) allow the mix of permanent and temporary labor to depend on the firm's productivity and the other state variables (through Δ_{jt}).

Our data combines the wages of permanent and temporary workers into $W_{jt} = W_{Pjt}(1-S_{Tjt}) + W_{Tjt}S_{Tjt}$, where $S_{Tjt} = \frac{L_{Tjt}}{L_{jt}}$ is the (quantity) share of temporary labor and $L_{jt} = L_{Pjt} + L_{Tjt}$ is hours worked by permanent and temporary workers in our data. To make do, we assume that the aggregator $\Lambda(L_{Pjt}, L_{Tjt})$ is linearly homogenous. This implies $L_{jt}^{*} = L_{jt}\Lambda(1-S_{Tjt}, S_{Tjt})$, $\frac{\partial L_{jt}^{*}}{\partial L_{Pjt}} = \Lambda_P(1-S_{Tjt}, S_{Tjt})$, and $\frac{\partial L_{jt}^{*}}{\partial L_{Tjt}} = \Lambda_T(1-S_{Tjt}, S_{Tjt})$. Using Euler's theorem to combine equations (10) and (11) yields

$$\begin{aligned}
&\nu\mu X_{jt}^{-\left(1+\frac{\nu\sigma}{1-\sigma}\right)} \exp\left(\omega_{Hjt}\right) \exp\left(-\frac{1-\sigma}{\sigma}\omega_{Ljt}\right) L_{jt}^{-\frac{1}{\sigma}} \Lambda(1-S_{Tjt}, S_{Tjt})^{-\frac{1-\sigma}{\sigma}} \\
&= \frac{W_{jt}\left(1+\frac{\Delta_{jt}}{1+\frac{W_{Tjt}}{W_{Pjt}}\frac{S_{Tjt}}{1-S_{Tjt}}}\right)}{P_{jt}\left(1-\frac{1}{\eta(p_{jt},D_{jt})}\right)} = \frac{W_{jt}\left(\frac{\frac{\Lambda_P(1-S_{Tjt},S_{Tjt})}{\Lambda_T(1-S_{Tjt},S_{Tjt})}+\frac{S_{Tjt}}{1-S_{Tjt}}}{\frac{W_{Pjt}}{W_{Tjt}}+\frac{S_{Tjt}}{1-S_{Tjt}}}\right)}{P_{jt}\left(1-\frac{1}{\eta(p_{jt},D_{jt})}\right)},
\end{aligned} \quad (12)$$

where the second equality follows from dividing equations (10) and (11) and solving for Δ_{jt}.

Because our data does not have the ratio $\frac{W_{Pjt}}{W_{Tjt}}$, we assume that $\frac{W_{Pjt}}{W_{Tjt}} = \lambda_0$ is an (unknown) constant[19] and treat $\frac{\frac{\Lambda_P(1-S_{Tjt},S_{Tjt})}{\Lambda_T(1-S_{Tjt},S_{Tjt})}+\frac{S_{Tjt}}{1-S_{Tjt}}}{\lambda_0+\frac{S_{Tjt}}{1-S_{Tjt}}} = \lambda_1(S_{Tjt})$ as an (unknown) function of S_{Tjt} that must be estimated nonparametrically along with the parameters of the production function. Because equation (12) presumes interior solutions for permanent and temporary

[19] In Appendix E, we use a wage regression to estimate wage premia of various types of labor. In the Online Appendix, we extend the specification and demonstrate that the wage premia do not change much if at all over time in line with our assumption that the ratio $\frac{W_{Pjt}}{W_{Ljt}}$ is constant.

labor, we exclude observations with $S_{Tjt} = 0$ and thus $L_{Tjt} = 0$ from the subsequent analysis.[20]

Turning from the labor to the materials decision, because the firm must maintain the ratio of outsourced to in-house materials Q_{Mjt}, the first-order condition for in-house materials is

$$\nu\beta_M \mu X_{jt}^{-\left(1+\frac{\nu\sigma}{1-\sigma}\right)} \exp\left(\omega_{Hjt}\right) \left(M_{jt}^*\right)^{-\frac{1}{\sigma}} \frac{dM_{jt}^*}{dM_{Ijt}} = \frac{P_{Ijt} + P_{Ojt}Q_{Mjt}}{P_{jt}\left(1 - \frac{1}{\eta(p_{jt}, D_{jt})}\right)}, \tag{13}$$

where $P_{Ijt} + P_{Ojt}Q_{Mjt}$ is the effective cost of an additional unit of in-house materials.

Our data has the materials bill $P_{Mjt}M_{jt} = P_{Ijt}M_{Ijt} + P_{Ojt}M_{Ojt}$, the (value) share of outsourced materials $S_{Ojt} = \frac{P_{Ojt}M_{Ojt}}{P_{Mjt}M_{jt}}$, and the price of materials P_{Mjt}. We assume $P_{Mjt} = P_{Ijt} + P_{Ojt}Q_{Mjt}$ so that the price of materials is the effective cost of an additional unit of in-house materials. This implies $M_{jt} = M_{Ijt}$. To map the model to the data, we further assume that $\Gamma(M_{Ijt}, M_{Ojt})$ is linearly homogenous and normalize $\Gamma(M_{Ijt}, 0) = M_{Ijt}$. This implies $M_{jt}^* = M_{Ijt}\Gamma\left(1, \frac{P_{Ijt}}{P_{Ojt}}\frac{S_{Ojt}}{1-S_{Ojt}}\right)$ and $\frac{dM_{jt}^*}{dM_{Ijt}} = \Gamma\left(1, \frac{P_{Ijt}}{P_{Ojt}}\frac{S_{Ojt}}{1-S_{Ojt}}\right)$. Rewriting equation (13) yields

$$\nu\beta_M \mu X_{jt}^{-\left(1+\frac{\nu\sigma}{1-\sigma}\right)} \exp\left(\omega_{Hjt}\right) M_{jt}^{-\frac{1}{\sigma}} \Gamma\left(1, \frac{P_{Ijt}}{P_{Ojt}}\frac{S_{Ojt}}{1-S_{Ojt}}\right)^{-\frac{1-\sigma}{\sigma}} = \frac{P_{Mjt}}{P_{jt}\left(1 - \frac{1}{\eta(p_{jt}, D_{jt})}\right)}. \tag{14}$$

Because our data does not have the ratio $\frac{P_{Ijt}}{P_{Ojt}}$, we assume that $\frac{P_{Ijt}}{P_{Ojt}} = \gamma_0$ is an (unknown) constant and treat $\ln\Gamma\left(1, \gamma_0\frac{S_{Ojt}}{1-S_{Ojt}}\right) = \gamma_1(S_{Ojt})$ as an (unknown) function of S_{Ojt}.[21] Equation (14) presumes an interior solution for in-house materials; it is consistent with a corner solution for outsourced materials. Indeed, absent outsourcing equation (14) reduces to the first-order condition for in-house materials.

Our primary interest is the bias of technological change. We thus think of $\lambda_1(S_{Tjt})$ and $\gamma_1(S_{Ojt})$ as "correction terms" on labor and, respectively, materials that help account for the substantial heterogeneity across the firms within an industry. Because we estimate these terms nonparametrically, they can accommodate different theories about the Spanish labor market and the role of outsourcing. For example, we develop an alternative model of outsourcing in the Online Appendix that assumes that both in-house and outsourced materials are static inputs that the firm may mix and match at will.

[20] Compare columns (1) and (2) of Tables 1 and 3 with columns (1) and (2) of Table 4 for the exact number of observations and firms we exclude.

[21] We have experimented with assuming that $\frac{P_{Ijt}}{P_{Ojt}} = \gamma_0(t)$ is an (unknown) function of time t and treating $\ln\Gamma\left(1, \gamma_0(t)\frac{S_{Ojt}}{1-S_{Ojt}}\right) = \gamma_1\left(\gamma_0(t)\frac{S_{Ojt}}{1-S_{Ojt}}\right)$ as an (unknown) function of $\gamma_0(t)S_{Ojt}$. As we show in the Online Appendix, not much changes. Equation (17) tends to yield somewhat lower estimates of σ compared to our leading estimates in column (3) of Table 4. Compared to our leading estimates in columns (1) and (2) of Table 7 equation (20) tends to yield somewhat lower estimates of β_K and similar estimates of ν in the eight industries where we have been able to obtain estimates. Our conclusions about the bias of technological change remain the same.

Labor-augmenting and Hicks-neutral productivity. From the labor and materials decisions in equations (12) and (14) we recover (conveniently rescaled) labor-augmenting productivity $\widetilde{\omega}_{Ljt} = (1-\sigma)\omega_{Ljt}$ and Hicks-neutral productivity ω_{Hjt} as

$$\begin{aligned} \widetilde{\omega}_{Ljt} &= \widetilde{\gamma}_L + m_{jt} - l_{jt} + \sigma(p_{Mjt} - w_{jt}) - \sigma\lambda_2(S_{Tjt}) + (1-\sigma)\gamma_1(S_{Ojt}) \\ &\equiv \widetilde{h}_L(m_{jt} - l_{jt}, p_{Mjt} - w_{jt}, S_{Tjt}, S_{Ojt}), \qquad (15) \\ \omega_{Hjt} &= \gamma_H + \frac{1}{\sigma} m_{jt} + p_{Mjt} - p_{jt} - \ln\left(1 - \frac{1}{\eta(p_{jt}, D_{jt})}\right) \\ &\quad + \left(1 + \frac{\nu\sigma}{1-\sigma}\right) x_{jt} + \frac{1-\sigma}{\sigma}\gamma_1(S_{Ojt}) \\ &\equiv h_H(k_{jt}, m_{jt}, S_{Mjt}, p_{jt}, p_{Mjt}, D_{jt}, S_{Tjt}, S_{Ojt}), \qquad (16) \end{aligned}$$

where $\widetilde{\gamma}_L = -\sigma \ln \beta_M$, $\lambda_2(S_{Tjt}) = \ln\left(\lambda_1(S_{Tjt})\Lambda\left(1 - S_{Tjt}, S_{Tjt}\right)^{\frac{1-\sigma}{\sigma}}\right)$, $\gamma_H = -\ln(\nu\beta_M\mu)$,

$$X_{jt} = \beta_K K_{jt}^{-\frac{1-\sigma}{\sigma}} + \beta_M \left(M_{jt} \exp\left(\gamma_1(S_{Ojt})\right)\right)^{-\frac{1-\sigma}{\sigma}} \left(\frac{1 - S_{Mjt}}{S_{Mjt}}\lambda_1(S_{Tjt}) + 1\right),$$

and $S_{Mjt} = \frac{P_{Mjt}M_{jt}}{VC_{jt}}$ is the share of materials in variable cost $VC_{jt} = W_{jt}L_{jt} + P_{Mjt}M_{jt}$. Recall that uppercase letters denote levels and lowercase letters denote logs. The functions $\widetilde{h}_L(\cdot)$ and $h_H(\cdot)$ allow us to recover unobservable labor-augmenting productivity $\widetilde{\omega}_{Ljt}$ and Hicks-neutral productivity ω_{Hjt} from observables, and we refer to them as inverse functions from here on. Without loss of generality, we set $\beta_K + \beta_M = 1$ in what follows.

5 Empirical strategy

The endogeneity problem in production function estimation arises because a firm's decisions depend on its productivity, and productivity is not observed by the econometrician. However, if the firm's productivity can be inferred from its decisions, then it can be controlled for in the estimation. To do so, we combine the inverse functions in equations (15) and (16) with the laws of motion for labor-augmenting and Hicks-neutral productivity in equations (7) and (8) into estimation equations for the parameters of the production function in equation (6).

Labor-augmenting productivity. We use equation (15) to recover labor-augmenting productivity $\widetilde{\omega}_{Ljt}$ and equation (7) to model its evolution. Substituting the inverse function in equation (15) into the law of motion in equation (7), we form our first estimation equation

$$\begin{aligned} m_{jt} - l_{jt} = -\sigma(p_{Mjt} - w_{jt}) + \sigma\lambda_2(S_{Tjt}) - (1-\sigma)\gamma_1(S_{Ojt}) \\ + \widetilde{g}_{Lt-1}(\widetilde{h}_L(m_{jt-1} - l_{jt-1}, p_{Mjt-1} - w_{jt-1}, S_{Tjt-1}, S_{Ojt-1}), R_{jt-1}) + \widetilde{\xi}_{Ljt}, \qquad (17) \end{aligned}$$

where the (conveniently rescaled) conditional expectation function is

$$\widetilde{g}_{Lt-1}(\widetilde{h}_L(\cdot), R_{jt-1}) = (1-\sigma) g_{Lt-1}\left(\frac{\widetilde{h}_L(\cdot)}{1-\sigma}, R_{jt-1}\right)$$

and $\widetilde{\xi}_{Ljt} = (1-\sigma)\xi_{Ljt}$.[22]

We allow $\widetilde{g}_{Lt-1}(\widetilde{h}_L(\cdot), R_{jt-1})$ to differ between zero and positive R&D expenditures and specify

$$\begin{aligned} \widetilde{g}_{Lt-1}(\widetilde{h}_L(\cdot), R_{t-1}) \quad = \quad & \widetilde{g}_{L0}(t-1) + 1(R_{jt-1} = 0)\widetilde{g}_{L1}(\widetilde{h}_L(\cdot)) \\ & +1(R_{jt-1} > 0)\widetilde{g}_{L2}(\widetilde{h}_L(\cdot), r_{jt-1}), \end{aligned} \tag{18}$$

where $1(\cdot)$ is the indicator function and the functions $\widetilde{g}_{L1}(\widetilde{h}_L(\cdot))$ and $\widetilde{g}_{L2}(\widetilde{h}_L(\cdot), r_{jt-1})$ are modeled as described in Appendix C. Because the Markov process governing labor-augmenting productivity is time-inhomogeneous, we allow the conditional expectation function $\widetilde{g}_{Lt-1}(\widetilde{h}_L(\cdot), R_{jt-1})$ to shift over time by $\widetilde{g}_{L0}(t-1)$. In practice, we model this shift with time dummies.

Compared to directly estimating equation (5) by OLS, equation (17) intuitively diminishes the endogeneity problem because breaking out the part of $\widetilde{\omega}_{Ljt}$ that is observable via the conditional expectation function $\widetilde{g}_{Lt-1}(\cdot)$ leaves "less" in the error term. This also facilitates instrumenting for any remaining correlation between the included variables and the error term.

In our model, labor l_{jt}, materials m_{jt}, the wage w_{jt}, and the share of temporary labor S_{Tjt} are correlated with the productivity innovation $\widetilde{\xi}_{Ljt}$ (since $\widetilde{\xi}_{Ljt}$ is part of $\widetilde{\omega}_{Ljt}$). Note that $w_{jt} = \ln(W_{Pjt}(1 - S_{Tjt}) + W_{Tjt}S_{Tjt})$ may be correlated with $\widetilde{\xi}_{Ljt}$ even though the firm takes the wage of permanent workers W_{Pjt} and the wage of temporary workers W_{Tjt} as given because S_{Tjt} may depend on $\widetilde{\omega}_{Ljt}$ and ω_{Hjt} through equations (10) and (11). We therefore base estimation on the moment conditions

$$E\left[A_{Ljt}(z_{jt})\widetilde{\xi}_{Ljt}\right] = 0, \tag{19}$$

where $A_{Ljt}(z_{jt})$ is a vector of functions of the exogenous variables z_{jt} as described in Appendix C.

In considering instruments it is important to keep in mind that equation (17) models the evolution of labor-augmenting productivity $\widetilde{\omega}_{Ljt}$. As a consequence, instruments have to be uncorrelated with the productivity innovation $\widetilde{\xi}_{Ljt}$ but not necessarily with productivity itself. Because $\widetilde{\xi}_{Ljt}$ is the innovation to productivity $\widetilde{\omega}_{Ljt}$ in period t, it is not known to the

[22] Equation (17) is a semiparametric, partially linear, model with the additional restriction that the inverse function $\widetilde{h}_L(\cdot)$ is of known form. Identification in the sense of the ability to separate the parametric and nonparametric parts of the model follows from standard arguments (Robinson 1988, Newey, Powell & Vella 1999).

firm when it makes its decisions in period $t-1$. All past decisions are therefore uncorrelated with $\widetilde{\xi}_{Ljt}$. In particular, having been decided in period $t-1$, l_{jt-1} and m_{jt-1} are uncorrelated with $\widetilde{\xi}_{Ljt}$, although they are correlated with $\widetilde{\omega}_{Ljt}$ as long as productivity is correlated over time. Similarly, because S_{Tjt-1} and thus $w_{jt-1} = \ln\left(W_{Pjt-1}(1 - S_{Tjt-1}) + W_{Tjt-1}S_{Tjt-1}\right)$ are determined in period $t-1$, they are uncorrelated with the productivity innovation $\widetilde{\xi}_{Ljt}$ in period t. We therefore use lagged labor l_{jt-1}, lagged materials m_{jt-1}, and the lagged wage w_{jt-1} for instruments.

In contrast to the wage w_{jt}, in our model the price of materials $p_{Mjt} = \ln\left(P_{Ijt} + P_{Ojt}Q_{Mjt}\right)$ is uncorrelated with $\widetilde{\xi}_{Ljt}$ because the ratio of outsourced to in-house materials Q_{Mjt} is determined in period $t-1$. For the same reason, the share of outsourced materials $S_{Ojt} = \frac{P_{Ojt}Q_{Mjt}}{P_{Ijt}+P_{Ojt}Q_{Mjt}}$ is uncorrelated with $\widetilde{\xi}_{Ljt}$. We nevertheless choose to err on the side of caution and restrict ourselves to the lagged price of materials p_{Mjt-1} and the lagged share of outsourcing S_{Ojt-1} for instrument. Finally, time t and the demand shifter D_{jt} are exogenous by construction and we use them for instruments.

The reasoning that the timing of decisions and the Markovian assumption on the evolution of productivity taken together imply that all past decisions are uncorrelated with productivity innovations originates in Olley & Pakes (1996). The subsequent literature uses it to justify lagged input quantities as instruments (see, e.g., Section 2.4.1 of Ackerberg et al. (2007)). In Doraszelski & Jaumandreu (2013), we extend this reasoning to justify lagged output and input prices as instruments (pp. 1347–1348). More recently, De Loecker, Goldberg, Khandelwal & Pavcnik (2016) do the same to justify the lagged price of output as instrument (p. 471).

A test for overidentifying restrictions in Section 6 cannot reject the validity of the moment conditions in equation (19). As discussed there, this is in large part because the aggregators $\Lambda(L_{Pjt}, L_{Tjt})$ and $\Gamma(M_{Ijt}, M_{Ojt})$ and the correction terms $\lambda_2(S_{Tjt})$ and $\gamma_1(S_{Ojt})$ associated with them account for quality differences between permanent and temporary labor, respectively, in-house and outsourced materials and differences in the use of these inputs over time and across firms. Absent these correction terms, one may be concerned that unobserved quality lingers in quantities and, perhaps even more so, in prices. This may lead to correlation between the lagged input prices w_{jt-1} and p_{Mjt-1} and the productivity innovation $\widetilde{\xi}_{Ljt}$ and invalidate them as instruments. By controlling for the composition of inputs, the correction terms in equation (17) absorb quality differences.[23] Indeed, additional checks suggest that there is limited reason to doubt that w_{jt-1} and p_{Mjt-1} are uncorrelated with $\widetilde{\xi}_{Ljt}$, in line with the test for overidentifying restrictions.

To the extent that a concern remains, it must thus draw on the notion that quality differences at a finer level play an important role. We address this concern in two ways by leveraging our data on the skill mix of a firm's labor force. First, in our data the larger

[23] De Loecker et al. (2016) develop a model that links the quality of a firm's inputs with the quality of its output. This allows them to use the the observed output price to control for unobserved input prices and quality. Depending on data availability, this may be an alternative to our approach.

part of the variation in the wage across firms and periods can be attributed to geographic and temporal differences in the supply of labor and the fact that firms operate in different product submarkets (see Appendix E). This part of the variation is arguably exogenous and therefore useful for estimating equation (17). The smaller part of the variation in the wage can be attributed to differences in the skill mix and the quality of labor that may potentially be correlated with the error term in equation (17).[24] However, we show in Section 6 that our estimates are robust to purging the variation due to differences in the skill mix from the lagged wage w_{jt-1}. Second, in Section 6 we explicitly model quality differences at a finer level by assuming that the firm faces a menu of qualities and wages in the market for permanent labor.

Hicks-neutral productivity. Substituting the inverse functions in equations (15) and (16) into the production function in equation (6) and the law of motion for Hicks-neutral productivity ω_{Hjt} in equation (8), we form our second estimation equation[25,26]

$$y_{jt} = -\frac{\nu\sigma}{1-\sigma}x_{jt}$$
$$+g_{Ht-1}(h_H(k_{jt-1}, m_{jt-1}, S_{Mjt-1}, p_{jt-1}, p_{Mjt-1}, D_{jt-1}, S_{Tjt-1}, S_{Ojt-1}), R_{jt-1}) + \xi_{Hjt} + e_{jt}. \tag{20}$$

We specify $g_{Ht-1}(h_H(\cdot), R_{jt-1})$ analogously to $\widetilde{g}_{Lt-1}(\widetilde{h}_L(\cdot), R_{jt-1})$ in equation (18).

Because output y_{jt}, materials m_{jt}, the share of materials in variable cost S_{Mjt}, and the share of temporary labor S_{Tjt} are correlated with ξ_{Hjt} in our model, we base estimation on the moment conditions

$$E\left[A_{Hjt}(z_{jt})(\xi_{Hjt} + e_{jt})\right] = 0,$$

where $A_{Hjt}(z_{jt})$ is a vector of function of the exogenous variables z_{jt}. As before, we exploit the timing of decisions and the Markovian assumption on the evolution of productivity to rely on lags for instruments. In addition, $k_{jt} = \ln((1-\delta)K_{jt-1} + I_{jt-1})$ is determined in period $t-1$ and therefore uncorrelated with ξ_{Hjt}.

Estimation. We use the two-step GMM estimator of Hansen (1982). Let $\nu_{Ljt}(\theta_L) = \widetilde{\xi}_{Ljt}$ be the residual of estimation equation (17) as a function of the parameters θ_L to be estimated

[24] A parallel discussion applies to materials. Kugler & Verhoogen (2012) point to differences in the quality of materials whereas Atalay (2014) documents substantial variation in the price of materials across plants in narrowly defined industries with negligible quality differences. This variation is partly due to geography and differences in cost and markup across suppliers that are arguably exogenous to a plant.

[25] There are other possible estimation equations. In particular, one can use the labor and materials decisions in equations (12) and (14) together with the production function in equation (6) to recover $\widetilde{\omega}_{Ljt}$, ω_{Hjt}, and e_{jt} and then set up separate moment conditions in $\widetilde{\xi}_{Ljt}$, ξ_{Hjt}, and e_{jt}. This may yield efficiency gains. Our estimation equation (20) has the advantage that it is similar to a CES production function that has been widely estimated in the literature.

[26] Equation (20) is again a semiparametric model with the additional restriction that the inverse function $h_H(\cdot)$ is of known form.

and $\nu_{Hjt}(\theta_H) = \xi_{Hjt} + e_{jt}$ the residual of estimation equation (20) as a function of θ_H. The GMM problem corresponding to equation (17) is

$$\min_{\theta_L} \left[\frac{1}{N} \sum_j A_{Lj}(z_j) \nu_{Lj}(\theta_L) \right]' \widehat{W}_L \left[\frac{1}{N} \sum_j A_{Lj}(z_j) \nu_{Lj}(\theta_L) \right], \tag{21}$$

where $A_{Lj}(z_j)$ is a $Q_L \times T_j$ matrix of functions of the exogenous variables z_j, $\nu_{Lj}(\theta_L)$ is a $T_j \times 1$ vector, $\widehat{W}_L$ is a $Q_L \times Q_L$ weighting matrix, Q_L is the number of instruments, T_j is the number of observations of firm j, and N is the number of firms. We provide further details in Appendix C.

The GMM problem corresponding to equation (20) is analogous. Equation (20) is considerably more nonlinear than equation (17). To facilitate its estimation, we impose the estimated values of those parameters in θ_L that also appear in θ_H. We correct the standard errors as described in the Online Appendix. Because they tend to be more stable, we report first-step estimates for equation (20) and use them in the subsequent analysis; however, we use second-step estimates for testing.

6 Labor-augmenting technological change

From equation (17) we obtain an estimate of the elasticity of substitution and recover labor-augmenting productivity at the firm level.

Elasticity of substitution. Tables 3 and 4 summarize different estimates of the elasticity of substitution. To facilitate the comparison with the existing literature, we begin by proxying for $\widetilde{\omega}_{Ljt} = (1-\sigma)\omega_{Ljt}$ in equation (5) by a time trend $\widetilde{\delta}_L t$ and estimate by OLS. As can be seen from columns (3) and (4) of Table 3, with the exception of industry 9, the estimates of the elasticity of substitution are in excess of one, whereas the estimates in the previous literature lie somewhere between 0 and 1 (Chirinko 2008, Bruno 1984, Rotemberg & Woodford 1996, Oberfield & Raval 2014). This reflects, first, that a time trend is a poor proxy for labor-augmenting technological change at the firm level and, second, that the estimates are upward biased as a result of the endogeneity problem.

We address the endogeneity problem by modeling the evolution of labor-augmenting productivity and estimating equation (17) by GMM. To illustrate the role of controlling for the composition of inputs in our empirical strategy, it is helpful to abstract from the distinction between permanent and temporary labor and in-house and outsourced materials. To this end, we revert to the setting in Section 2 and assume that labor l_{jt} and materials m_{jt} are homogenous inputs that are chosen each period to maximize short-run profits. This implies $\lambda_1(S_{Tjt}) = 1$, $\lambda_2(S_{Tjt}) = 0$, and $\gamma_1(S_{Ojt}) = 0$, so that the correction terms on labor and materials vanish and equation (15) reduces to equation (5). Columns (5)–(10) of Table 3 refer to this simplified model. As expected the estimates of the elasticity of substitution are much lower and range from 0.45 to 0.64, as can be seen from column (5). With the exception of industries 6 and 8 in which σ is either implausibly high or low, we clearly reject the special cases of both a Leontieff ($\sigma \to 0$) and a Cobb-Douglas ($\sigma = 1$) production function.

Testing for overidentifying restrictions, however, we reject the validity of the moment conditions in the simplified model at a 5% level in five industries and we are close to rejecting in two more industries (columns (6) and (7)). To pinpoint the source of this problem, we exclude the subset of moments involving lagged materials m_{jt-1} from the estimation. As can be seen from columns (8)–(10), the resulting estimates of the elasticity of substitution lie between 0.46 and 0.85 in all industries and at a 5% level we can no longer reject the validity of the moment conditions in any industry.

To see why the exogeneity of lagged materials m_{jt-1} is violated contrary to the timing of decisions in our model, recall that a firm engages in outsourcing if it can procure customized parts and pieces from its suppliers that are cheaper or better than what the firm can make in house from scratch. Lumping in-house and outsourced materials together pushes these quality differences into the error term. As outsourcing often relies on contractual relationships between the firm and its suppliers, the error term is likely correlated over time and thus with lagged materials m_{jt-1} as well.

The correction term $\gamma_1(S_{Ojt})$ in equation (17) absorbs quality differences between in-house and outsourced materials into the aggregator $\Gamma(M_{Ijt}, M_{Ojt})$ and accounts for the wedge that outsourcing may drive between the relative quantities and prices of materials and labor. The correction term $\lambda_2(S_{Tjt})$ similarly absorbs quality differences between permanent and temporary labor into the aggregator $\Lambda(L_{Pjt}, L_{Tjt})$ and accounts for adjustment costs on permanent labor. As can be seen in columns (3)–(5) of Table 4, the correction terms duly restore the exogeneity of lagged materials m_{jt-1} as we cannot reject the validity of the moment conditions at a 5% level in any industry except for industry 7 in which we (barely) reject.[27] Our leading estimates of σ in column (3) of Table 4 lie between 0.44 and 0.80. Compared to the estimates in column (8) of Table 3, there are no systematic changes and our leading estimates are somewhat lower in five industries and somewhat higher in five industries. In sum, accounting for outsourcing and adjustment costs on permanent labor is an improvement over the assumption in Levinsohn & Petrin (2003) and many others that labor and materials are homogenous and static inputs and a key step in estimating the elasticity of substitution.

Additional checks: Lagged input prices. To uncover any potential weaknesses of our leading specification, we supplement the omnibus test for overidentifying restrictions with additional checks. Because the lagged wage w_{jt-1} and the lagged price of materials

[27]As noted in Section 4, we exclude observations with $S_{Tjt} = 0$ and thus $L_{Tjt} = 0$ because equation (12) presumes interior solutions for permanent and temporary labor. Compare columns (1) and (2) of Tables 1 and 3 with columns (1) and (2) of Table 4 for the exact number of observations and firms we exclude.

p_{Mjt-1} play a key role in the estimation of equation (17), we conduct two Sargan difference tests to more explicitly validate their use as instruments. In case of w_{jt-1}, we compute the difference in the value of the GMM objective function when we exclude the subset of moments involving p_{Mjt-1} and when we exclude the subset of moments involving w_{jt-1} and p_{Mjt-1}; in case of p_{Mjt-1} we proceed analogously.[28] As can be seen in columns (6)–(9) of Table 4, the exogeneity assumption on the lagged wage is rejected at a 5% level in three industries, while that on the lagged price of materials cannot be rejected in any industry. Viewing all these tests in conjunction, to the extent that a concern about our leading specification is warranted, it appears more related to labor than to materials.

Below we use the available data on the skill mix to further model a firm facing a choice of different qualities and wages in the market for permanent labor. For now we note that our estimates of the elasticity of substitution are robust to purging the variation due to differences in the quality of labor from the lagged wage w_{jt-1}. In Appendix E, we use a wage regression to isolate the part of the wage that depends on the skill mix of a firm's labor force. Using $\widehat{w}_{Qjt-1}$ to denote this part, we replace w_{jt-1} as an instrument by $w_{jt-1}-\widehat{w}_{Qjt-1}$. Compared to column (3) of Table 4, the estimates of the elasticity of substitution in column (10) decrease somewhat in three industries, remain essentially unchanged in two industries, and increase somewhat in five industries.[29] The absence of substantial and systematic changes confirms that the variation in w_{jt-1} is exogenous with respect to $\widetilde{\xi}_{Ljt}$ and therefore useful in estimating equation (17), in line with the test for overidentifying restrictions.

Labor-augmenting technological change. With equation (17) estimated, we recover the labor-augmenting productivity $\omega_{Ljt} = \frac{\widetilde{\omega}_{Ljt}}{1-\sigma}$ of firm j in period t up to an additive constant from equation (15). In what follows, we therefore de-mean ω_{Ljt} by industry. Abusing notation, we continue to use ω_{Ljt} to denote the de-meaned labor-augmenting productivity of firm j in period t.

To obtain aggregate measures representing an industry, we account for the survey design by replicating the subsample of small firms $\frac{70\%}{5\%} = 14$ times before pooling it with the subsample of large firms. Unless noted otherwise, we report weighed averages of individual measures, where the weight $\mu_{jt} = P_{jt-2}Y_{jt-2}/\sum_j P_{jt-2}Y_{jt-2}$ is the share of sales of firm j in period $t-2$. Using the second lag reduces the covariance between the weight and the variable of interest and thus also the extent of reallocation in the sense of Olley & Pakes (1996).

The growth of labor-augmenting productivity at firm j in period t is $\Delta\omega_{Ljt} = \omega_{Ljt} -$

[28] To use the same weighting matrix for both specifications and not unduly change variances when we exclude subsets of moments, we delete the appropriate rows and columns from the weighting matrix for our leading specification.

[29] As we show in the Online Appendix, not much changes if we isolate the part of the wage that additionally depends on firm size to try and account for the quality of labor beyond our rather coarse data on the skill mix of a firm's labor force (Oi & Idson 1999). Compared to column (3) of Table 4, the estimates of the elasticity of substitution decrease somewhat in three industries, remain essentially unchanged in three industries, and increase somewhat in four industries.

ω_{Ljt-1}.[30] In line with the patterns in the data described in Section 3, our estimates imply an important role for labor-augmenting technological change. As can be seen from column (1) of Table 5, labor-augmenting productivity grows quickly, on average, with rates of growth ranging from 0.9% and 1.7% per year in industries 10 and 7 to 14.5% and 14.6% in industries 6 and 2 and above in industry 5.

Ceteris paribus $\Delta\omega_{Ljt} \approx \frac{\exp(\omega_{Ljt})L^*_{jt-1} - \exp(\omega_{Ljt-1})L^*_{jt-1}}{\exp(\omega_{Ljt-1})L^*_{jt-1}}$ approximates the rate of growth of a firm's effective labor force $\exp(\omega_{Ljt-1})L^*_{jt-1}$. To facilitate comparing labor-augmenting to Hicks-neutral productivity, we approximate the rate of growth of the firm's output Y_{jt-1} by $\epsilon_{Ljt-2}\Delta\omega_{Ljt}$, where ϵ_{Ljt-2} is the elasticity of output with respect to the firm's effective labor force in period $t-2$ (see Appendix D).[31] This output effect, while on average close to zero in industry 9, ranges from 0.6% per year in industry 7 to 3.0%, 3.1%, and 3.2% in industries 6, 2, and 4, see column (2) of Table 5. Across industries, labor-augmenting technological change causes output to grow by 1.6% per year.

Figure 1 illustrates the magnitude of the output effect of labor-augmenting technological change and the heterogeneity in its impact across industries. The depicted index cumulates the year-to-year changes and is normalized to one in 1991. Technological change appears to have slowed in the 2000s compared to the 1990s: across industries, labor-augmenting technological change causes output to grow by 2.1% per year before 2000 and by 1.0% per year after 2000.

Dispersion and persistence. A substantial literature documents dispersion and persistence in productivity (see Bartelsman & Doms (2000) and Syverson (2011) and the references therein). To be able to compare labor-augmenting productivity to Hicks-neutral productivity, we focus on $\epsilon_{Ljt-2}\omega_{Ljt}$. Because ω_{Ljt} is de-meaned, $\epsilon_{Ljt-2}\omega_{Ljt}$ measures the labor-augmenting productivity of firm j in period t relative to the average productivity, suitably converted into output terms. We thus refer to $\epsilon_{Ljt-2}\omega_{Ljt}$ as labor-augmenting productivity in output terms in what follows.

We measure dispersion by the interquartile range of $\epsilon_{Ljt-2}\omega_{Ljt}$. As can be seen from column (3) of Table 5, the interquartile range is between 0.24 in industry 9 and 0.70 in industry 6. This is comparable to the existing literature.[32] Turning from dispersion to persistence, $\epsilon_{Ljt-2}\omega_{Ljt}$ is highly autocorrelated (column (4)), indicating that differences in labor-augmenting productivity between firms persist over time.

[30] Given the specification of $\widetilde{g}_{Lt-1}(\widetilde{h}_L(\cdot), R_{jt-1})$ in equation (18), we exclude observations where a firm switches from performing to not performing R&D or vice versa between periods $t-1$ and t from the subsequent analysis. We further exclude observations where a firm switches from zero to positive outsourcing or vice versa.

[31] Because ϵ_{Ljt} depends on ω_{Ljt} as can be seen from equation (29), $\Delta\omega_{Ljt}$ is systematically negatively correlated with ϵ_{Ljt} and systematically positively correlated with ϵ_{Ljt-1}. Using ϵ_{Ljt-2} drastically reduces the correlation between the constituent parts of the output effect of labor-augmenting technological change.

[32] For U.S. manufacturing industries, Syverson (2004) reports an interquartile range of log labor productivity of 0.66.

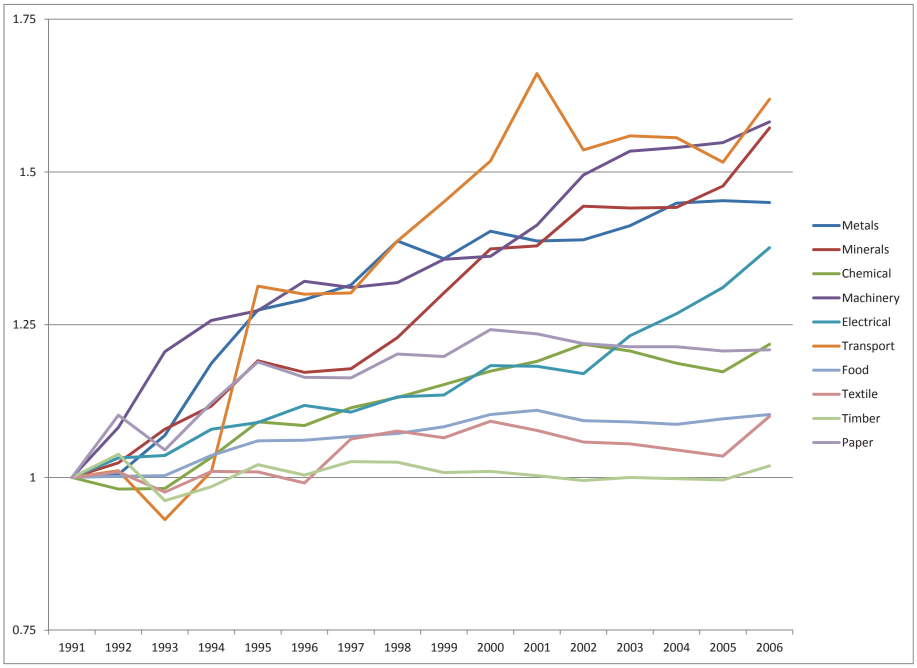

Figure 1: Output effect of labor-augmenting technological change. Index normalized to one in 1991.

Firms' R&D activities. As can be seen from column (5) of Table 5, firms that perform R&D have, on average, higher levels of labor-augmenting productivity in output terms than firms that do not perform R&D in all industries. In seven industries the output effect of labor-augmenting technological change for firms that perform R&D, on average, exceeds that of firms that do not perform R&D (columns (6) and (7)). Overall, our estimates indicate that firms' R&D activities are associated not only with higher levels of labor-augmenting productivity but by and large also with higher rates of growth of labor-augmenting productivity. Firms' R&D activities play a key role in determining the differences in labor-augmenting productivity across firms and the evolution of this dimension of productivity over time.

Firm turnover. To assess the impact of firm turnover on the output effect of labor-augmenting technological change, we classify a firm as a survivor if it enters the industry in or before 1990 and does not exit in or before 2006, as an exitor if it enters the industry in or before 1990 and exits in or before 2006, and as an entrant otherwise. Survivors account for most of the output effect of labor-augmenting technological change. Their contribution is 80% in industry 6 and above, except for industry 3 where the contribution of entrants is on par with the contribution of survivors. In the remaining industries, the contribution of entrants is small. The contribution of exitors is small in all industries.

Skill upgrading. In our data, there is a shift from unskilled to skilled workers. For example, the share of engineers and technicians in the labor force increases from 7.2% in 1991 to 12.3% in 2006. While this shift has to be seen against the backdrop of a general increase of university graduates in Spain during the 1990s and 2000s, it begs the question how much skill upgrading contributes to the growth of labor-augmenting productivity.

To answer this question—and to further alleviate concerns about the quality and composition of labor—we leverage our rather coarse data on the skill mix of a firm's labor force. Besides the share of temporary labor S_{Tjt}, our data has the share of white collar workers and the shares of engineers and technicians, respectively.[33]

We assume that there are Q types of permanent labor with qualities $1, \theta_2, \ldots, \theta_Q$ and corresponding wages $W_{P1jt}, W_{P2jt}, \ldots, W_{PQjt}$. The firm, facing this menu of qualities and wages, behaves as a price-taker in the labor market. In recognition of their different qualities, $L^*_{Pjt} = L_{P1jt} + \sum_{q=2}^{Q} \theta_q L_{Pqjt}$ is an aggregate of the Q types of permanent labor, with L_{Pqjt} being the quantity of permanent labor of type q at firm j in period t. $L^*_{jt} = \Lambda(L^*_{Pjt}, L_{Tjt})$ is the aggregate of permanent labor L^*_{Pjt} (instead of $L_{Pjt} = \sum_{q=1}^{Q} L_{Pqjt}$) and temporary labor L_{Tjt} in the production function in equation (6). Permanent labor is subject to convex adjustment costs $C_{B_P}(B_{Pjt}, B_{Pjt-1})$, where $B_{Pjt} = \sum_{q=1}^{Q} W_{Pqjt} L_{Pqjt}$ is the wage bill for permanent labor. The state vector Ω_{jt} therefore includes $B_{Pjt-1}, W_{P1jt}, W_{P2jt}, \ldots, W_{PQjt}$ instead of L_{Pjt-1} and W_{Pjt}.

The first-order condition for permanent labor of type q is

$$\nu\mu X_{jt}^{-\left(1+\frac{\nu\sigma}{1-\sigma}\right)} \exp\left(\omega_{Hjt}\right) \exp\left(-\frac{1-\sigma}{\sigma}\omega_{Ljt}\right) \left(L^*_{jt}\right)^{-\frac{1}{\sigma}} \frac{\partial L^*_{jt}}{\partial L^*_{Pjt}} \theta_q = \frac{W_{Pqjt}(1+\Delta_{jt})}{P_{jt}\left(1-\frac{1}{\eta(p_{jt},D_{jt})}\right)}, \quad (22)$$

where $\theta_1 = 1$ and the gap between the wage W_{Pqjt} and the shadow wage is

$$\begin{aligned} \Delta_{jt} &= \frac{\partial C_{B_P}(B_{Pjt}, B_{Pjt-1})}{\partial B_{Pjt}} - \frac{1}{W_{Pqjt}} \frac{1}{1+\rho} E_t\left[\frac{\partial V_{t+1}(\Omega_{jt+1})}{\partial L_{Pqjt}} | \Omega_{jt}, R_{jt}\right] \\ &= \frac{\partial C_{B_P}(B_{Pjt}, B_{Pjt-1})}{\partial B_{Pjt}} + \frac{1}{1+\rho} E_t\left[\frac{\partial C_{B_P}(B_{Pjt+1}, B_{Pjt})}{\partial B_{Pjt}} | \Omega_{jt}, R_{jt}\right]. \end{aligned}$$

Equation (22) implies that $\theta_q = \frac{W_{Pqjt}}{W_{P1jt}}$ at an interior solution. While our data does not have $W_{P1jt}, W_{P2jt}, \ldots, W_{PQjt}$, the wage regression in Appendix E enables us to recover θ_q by estimating the wage premium $\left(\frac{W_{Pqjt}}{W_{P1jt}} - 1\right)$ of permanent labor of type q over type 1.

Multiplying equation (22) by the share S_{Pqjt} of permanent workers of type q and sum-

[33] We have these latter measures in the year a firm enters the sample and every subsequent four years. We take the skill mix to be unchanging in the interim.

ming yields

$$\nu\mu X_{jt}^{-\left(1+\frac{\nu\sigma}{1-\sigma}\right)} \exp\left(\omega_{Hjt}\right) \exp\left(-\frac{1-\sigma}{\sigma}\omega_{Ljt}\right) \left(L_{jt}^{*}\right)^{-\frac{1}{\sigma}} \frac{\partial L_{jt}^{*}}{\partial L_{Pjt}^{*}}\Theta_{jt} = \frac{W_{Pjt}(1+\Delta_{jt})}{P_{jt}\left(1-\frac{1}{\eta(p_{jt},D_{jt})}\right)}, \tag{23}$$

where $\Theta_{jt} = S_{P1jt} + \sum_{q=2}^{Q}\theta_q S_{Pqjt} = 1 + \sum_{q=2}^{Q}\left(\frac{W_{Pqjt}}{W_{P1jt}} - 1\right) S_{Pqjt}$ is a quality index and $W_{Pjt} = \sum_{q=1}^{Q} W_{Pqjt}S_{Pqjt}$. Using Euler's theorem to combine equations (11) and (23) yields

$$\begin{aligned} &\nu\mu X_{jt}^{-\left(1+\frac{\nu\sigma}{1-\sigma}\right)} \exp\left(\omega_{Hjt}\right) \exp\left(-\frac{1-\sigma}{\sigma}\omega_{Ljt}\right) L_{jt}^{-\frac{1}{\sigma}} \Lambda((1-S_{Tjt})\Theta_{jt}, S_{Tjt})^{-\frac{1-\sigma}{\sigma}} \\ &= \frac{W_{jt}\left(1+\frac{\Delta_{jt}}{1+\frac{W_{Tjt}}{W_{Pjt}}\frac{S_{Tjt}}{1-S_{Tjt}}}\right)}{P_{jt}\left(1-\frac{1}{\eta(p_{jt},D_{jt})}\right)} = \frac{W_{jt}\left(\frac{\frac{\Lambda_P((1-S_{Tjt})\Theta_{jt},S_{Tjt})\Theta_{jt}}{\Lambda_T((1-S_{Tjt})\Theta_{jt},S_{Tjt})}+\frac{S_{Tjt}}{1-S_{Tjt}}}{\frac{W_{Pjt}}{W_{Tjt}}+\frac{S_{Tjt}}{1-S_{Tjt}}}\right)}{P_{jt}\left(1-\frac{1}{\eta(p_{jt},D_{jt})}\right)}, \end{aligned} \tag{24}$$

where the second equality follows from dividing equations (11) and (23) and solving for Δ_{jt}. We proceed as before by assuming that $\frac{W_{Pjt}}{W_{Ljt}} = \lambda_0$ is an (unknown) constant and treating $\frac{\frac{\Lambda_P((1-S_{Tjt})\Theta_{jt},S_{Tjt})\Theta_{jt}}{\Lambda_T((1-S_{Tjt})\Theta_{jt},S_{Tjt})}+\frac{S_{Tjt}}{1-S_{Tjt}}}{\lambda_0+\frac{S_{Tjt}}{1-S_{Tjt}}} = \lambda_1(S_{Tjt},\Theta_{jt})$ as an (unknown) function of S_{Tjt} and Θ_{jt} that must be estimated nonparametrically. Replacing $\lambda_2(S_{Tjt}) = \ln\left(\lambda_1(S_{Tjt})\Lambda(1-S_{Tjt},S_{Tjt})^{\frac{1-\sigma}{\sigma}}\right)$ by $\lambda_2(S_{Tjt},\Theta_{jt}) = \ln\left(\lambda_1(S_{Tjt},\Theta_{jt})\Lambda((1-S_{Tjt})\Theta_{jt},S_{Tjt})^{\frac{1-\sigma}{\sigma}}\right)$ in our estimation equation (17) therefore accounts for types of permanent labor that differ in their qualities and wages.

The estimates of the elasticity of substitution in column (8) of Table 5 continue to hover around 0.6 across industries, with the exception of industries 4 and 8 in which they are implausibly low. Compared to column (3) of Table 4, they decrease somewhat in three industries, remain essentially unchanged in two industries, and increase somewhat in five industries. This further supports the notion that quality differences at a finer level than permanent and temporary labor are of secondary importance for estimating equation (17).

We develop the quality index Θ_{jt} mainly to "chip away" at the productivity residual by improving the measurement of inputs in the spirit of the productivity literature (Griliches 1964, Griliches & Jorgenson 1967, Jorgenson 1995*a*, Jorgenson 1995*b*). As can be seen from column (11) of Table 5, skill upgrading indeed explains some, but by no means all of the growth of labor augmenting productivity. Compared to column (1), the rates of growth stay the same or go down in all industries. In industries 7, 8, 9, and 10 labor-augmenting productivity is stagnant or declining after accounting for skill upgrading, indicating that improvements in the skill mix over time are responsible for most of the growth of labor-augmenting productivity. In contrast, in industries 1, 2, 3, 4, 5, and 6, labor-augmenting productivity continues to grow after accounting for skill upgrading, albeit often at a much slower rate. In these industries, labor-augmenting productivity grows also because workers with a given set of skills become more productive over time.

7 The decline of the aggregate share of labor

In many advanced economies the aggregate share of labor in income has declined in past decades. While this decline has attracted considerable attention in the academic literature (Blanchard 1997, Bentolila & Saint-Paul 2004, McAdam & Willman 2013, Karabarbounis & Neiman 2014, Oberfield & Raval 2014) and in the public discussion following Piketty (2014), its causes and consequences remain contested. We use our estimates to show that biased technological change is the primary driver of the decline of the aggregate share of labor in the Spanish manufacturing sector over our sample period.

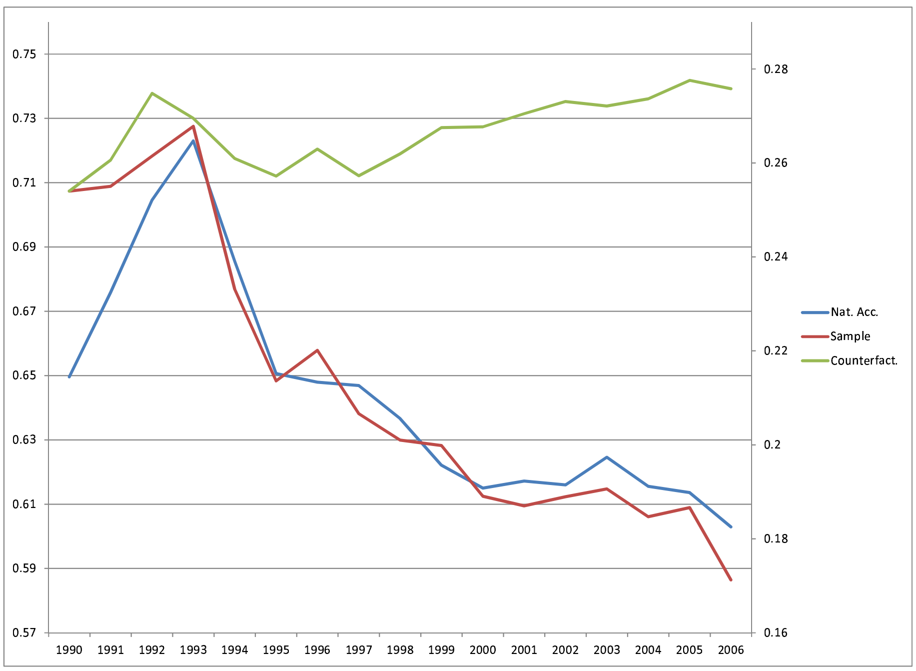

Figure 2: Aggregate share of labor in value added in National Accounts (left axis) and aggregate share of labor in variable cost in sample and counterfactual (right axis). Latter indices cumulate year-to-year changes using level in 1990 as base and average over industries using their share of total value added in column (4) of Table A1 as weight.

Let $VC_{Ljt} = W_{jt}L_{jt}$ be the wage bill, $VC_{jt} = W_{jt}L_{jt} + P_{Mjt}M_{jt}$ variable cost, and $S_{Ljt} = \frac{VC_{Ljt}}{VC_{jt}}$ the share of labor in variable cost of firm j in period t. Let $VC_{Lt} = \sum_j VC_{Ljt}$ and $VC_t = \sum_j VC_{jt}$ be the corresponding industry-wide aggregates. We focus on the aggregate share of labor in variable cost

$$S_{Lt} = \frac{VC_{Lt}}{VC_t} = \sum_j \frac{VC_{Ljt}}{VC_{jt}} \frac{VC_{jt}}{VC_t} = \sum_j S_{Ljt}\theta_{jt},$$

where $\theta_{jt} = \frac{VC_{jt}}{VC_t}$ is the variable cost of firm j in period t as a fraction of aggregate variable cost. As can be seen in Figure 2, the aggregate share of labor in variable cost closely tracks the aggregate share of labor in value added in the Spanish manufacturing sector in the National Accounts[34] over our sample period.

The year-to-year change in the aggregate share of labor in variable cost is $S_{Lt} - S_{Lt-1}$. Cumulated over our sample period, the decline of the aggregate share of labor ranges from 0.01 and 0.05 in industries 9 and 4 to 0.15 and 0.19 in industries 2 and 5, as can be seen in column (1) of Table 6.[35] To obtain insight into this decline, we build on Oberfield & Raval (2014) and decompose the year-to-year change as

$$S_{Lt} - S_{Lt-1} = \sum_j \theta_{jt}(S_{Ljt} - S_{Ljt-1}) + \sum_j (\theta_{jt} - \theta_{jt-1}) S_{Ljt-1}.$$

The second term captures reallocation across firms. Our model enables us to further decompose the first term. Rewriting equation (15) yields

$$S_{Ljt} = \frac{1}{1 + \exp\left(-\widetilde{\gamma}_L + (1-\sigma)\left(p_{Mjt} - w_{jt} + \omega_{Ljt}\right) + \sigma\lambda_2(S_{Tjt}) - (1-\sigma)\gamma_1(S_{Ojt})\right)}. \quad (25)$$

The first term may thus be driven by a change in the price of materials p_{Mjt} relative to the price of labor w_{jt}, a change in labor-augmenting productivity ω_{Ljt}, a change in the share of temporary labor S_{Tjt}, and a change in the share of outsourced materials S_{Ojt}. To quantify these drivers, we use a second-order approximation to $S_{Ljt} - S_{Ljt-1}$ as described in Appendix F.

We report the decomposition of the year-to-year change, cumulated over our sample period, in columns (2)–(7) of Table 6. The small size of the residual in column (7) indicates that our second-order approximation to $S_{Ljt} - S_{Ljt-1}$ readily accommodates nonlinearities. As can be seen in column (3), biased technological change emerges as the main force behind the decline of the aggregate share of labor. Changes in input prices in column (2) attenuate the decline. In contrast, the impact of temporary labor, outsourced materials, and reallocation across firms in the remaining columns is sometimes positive and sometimes negative and mostly small.

We use our model to compute the counterfactual evolution of the aggregate share of labor absent biased technological change by zeroing out the change in labor-augmenting productivity ω_{Ljt} in the decomposition of the year-to-year change. As can be seen in Figure 2, absent biased technological change the aggregate share of labor remains roughly constant over our sample period. We emphasize that this counterfactual holds fixed not only reallocation across firms but also the evolution of input prices, temporary labor, and outsourced materials. This may be questionable over longer stretches of time.

[34]Contabilidad Nacional de España, Bases 1986 and 1995, Instituto Nacional de Estadistica.

[35]We estimate $S_{Lt} - S_{Lt-1}$ as well as the various terms of the decomposition using firms that are in the sample in periods t and $t-1$.

Our conclusion that biased technological change is the primary driver of the decline of the aggregate share of labor echoes that of Oberfield & Raval (2014). Oberfield & Raval (2014) develop a decomposition of the change in the aggregate share of labor in value added in the U.S. manufacturing sector from 1970 to 2010. Perhaps the most important difference between their decomposition and ours is that we directly measure the bias of technological change at the level of the individual firm, whereas Oberfield & Raval (2014) treat it as the residual of their decomposition. Despite this difference and the different data sets used, the decompositions are complementary and both point to the overwhelming role of biased technological change in the decline of the aggregate share of labor.

8 Hicks-neutral technological change

From equation (17) we obtain an estimate of the elasticity of substitution and recover labor-augmenting productivity at the firm level. To recover Hicks-neutral productivity and the remaining parameters of the production function, we have to estimate equation (20).

Distributional parameters and elasticity of scale. Table 7 reports the distributional parameters β_K and $\beta_M = 1 - \beta_K$ and the elasticity of scale ν. Our estimates of β_K range from 0.07 in industry 8 to 0.31 in industry 6 (column (1)). Although the estimates of the elasticity of scale are rarely significantly different from one, taken together they suggest slightly decreasing returns to scale (columns (2)). We cannot reject the validity of the moment conditions in any industry by a wide margin (columns (3) and (4)).[36]

Price elasticity. Column (5) of Table 7 reports the average absolute value of the price elasticity $\eta(p_{jt-1}, D_{jt-1})$ implied by our estimates. It ranges from 1.79 in industry 9 to 6.04 and 9.11 in industries 5 and 2 and averages 3.20 across industries.[37]

Elasticity of substitution: Lagrange-multiplier test. The production function in equation (6) assumes that the elasticity of substitution between capital, labor, and materials is the same. We compare our leading specification to the more general nested CES production function

$$Y_{jt} = \left[\beta_K K_{jt}^{\frac{-(1-\tau)}{\tau}} + \left[\left(\exp(\omega_{Ljt}) L_{jt}^* \right)^{\frac{-(1-\sigma)}{\sigma}} + \beta_M \left(M_{jt}^* \right)^{\frac{-(1-\sigma)}{\sigma}} \right]^{\frac{-\sigma}{1-\sigma} \frac{-(1-\tau)}{\tau}} \right]^{\frac{-\nu\tau}{1-\tau}} \exp(\omega_{Hjt}) \exp(e_{jt}),$$

where the additional parameter τ is the elasticity of substitution between capital and labor, respectively, materials. We show in the Online Appendix that our first estimation equation

[36] In light of this wide margin, we do not further probe the validity of lagged prices as instruments.

[37] For U.S. manufacturing industries, Oberfield & Raval (2014) report price elasticities in a somewhat narrower range between 2.91 and 5.22 with a roughly comparable average of 3.91 across industries.

(17) remains unchanged and generalize our second estimation equation (20). This allows us to conduct a Lagrange-multiplier test for $\tau = \sigma$. As can be seen in columns (6) and (7) of Table 7, we cannot reject the validity of our leading specification in any industry.

Hicks-neutral technological change. With equation (20) estimated, we recover the Hicks-neutral productivity ω_{Hjt} of firm j in period t up to an additive constant from equation (16); in what follows, we use ω_{Hjt} to denote the de-meaned Hicks-neutral productivity. We proceed as before to obtain aggregate measures representing an industry.

The growth of Hicks-neutral productivity at firm j in period t is $\Delta\omega_{Hjt} = \omega_{Hjt} - \omega_{Hjt-1}$. Ceteris paribus $\Delta\omega_{Hjt} \approx \frac{X_{jt-1}^{-\frac{\nu\sigma}{1-\sigma}} \exp(\omega_{Hjt}) \exp(e_{jt-1}) - X_{jt-1}^{-\frac{\nu\sigma}{1-\sigma}} \exp(\omega_{Hjt-1}) \exp(e_{jt-1})}{X_{jt-1}^{-\frac{\nu\sigma}{1-\sigma}} \exp(\omega_{Hjt-1}) \exp(e_{jt-1})}$ approximates the rate of growth of a firm's output Y_{jt-1} and is therefore directly comparable to the output effect of labor-augmenting technological change. As can be seen from column (1) of Table 8, Hicks-neutral productivity grows quickly in five industries, with rates of growth ranging, on average, from 1.2% per year in industry 8 to 4.4% in industry 1. It grows much more slowly or barely at all in three industries, with rates of growth below 0.5% per year. While there is considerable heterogeneity in the rate of growth of Hicks-neutral productivity across industries, Hicks-neutral technological change causes output to grow by 1.4% per year.

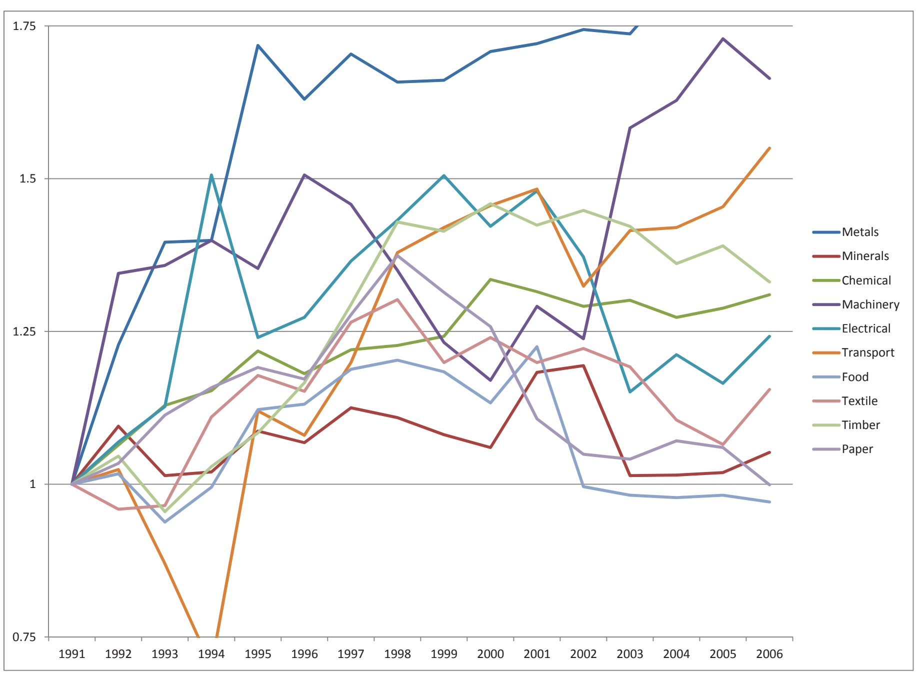

Figure 3: Hicks-neutral technological change. Index normalized to one in 1991.

Figure 3 illustrates the magnitude of Hicks-neutral technological change. The depicted index cumulates the year-to-year changes and is normalized to one in 1991.[38] The heterogeneity in the impact of Hicks-neutral technological change across industries clearly exceeds that of the output effect of labor-augmenting technological change (see again Figure 1). Once again, technological change appears to have slowed in the 2000s compared to the 1990s: across industries, Hicks-neutral technological change causes output to grow by 2.7% per year before 2000 and to shrink by 0.6% per year after 2000.

Dispersion and persistence. We measure dispersion by the interquartile range of ω_{Hjt}. As can be seen from column (2) of Table 8, the interquartile range is between 0.37 in industry 3 and 0.98 in industry 2.[39] Hicks-neutral productivity appears to be somewhat more disperse than labor-augmenting productivity in output terms. Once again, ω_{Hjt} is highly autocorrelated (column (3)), indicating that differences in Hicks-neutral productivity between firms persist over time.

Firms' R&D activities. As can be seen from column (4) of Table 8, firms that perform R&D have, on average, higher levels of Hicks-neutral productivity than firms that do not perform R&D in six industries but lower levels of Hicks-neutral productivity in four industries. While there is practically no difference in industry 10, the rate of growth of Hicks-neutral productivity for firms that perform R&D, on average, exceeds that of firms that do not perform R&D in five industries (columns (5) and (6)). Overall, our estimates indicate that firms' R&D activities are associated with higher levels and rates of growth of Hicks-neutral productivity, although firms' R&D activities seem less closely tied to Hicks-neutral than to labor-augmenting productivity. This is broadly consistent with the large literature on induced innovation that argues that firms direct their R&D activities to conserve the relatively more expensive factors of production, in particular labor.[40]

Firm turnover. Similar to the output effect of labor-augmenting technological change, survivors account for most of Hicks-neutral technological change. Their contribution is 62% in industry 9 and above. While the contributions of entrants and exitors are small in most industries, they are negative and more sizable in industries 2, 5, 7, 8, and 10. As a result, in these industries the rate of growth of Hicks-neutral productivity is 0.7%, 3.0%, 1.2%, 1.7%, and 1.2% amongst surviors compared to 0.5%, 2.0%, 0.1%, 1.2%, and 0.2% for all firms (see again column (1) of Table 8).

[38] In industry 9, in line with column (1) of Table 8, we trim values of $\Delta\omega_H$ below -0.25 and above 0.5.

[39] For Chinese manufacturing industries, Hsieh & Klenow (2009) report an interquartile range of log total factor productivity of 1.28.

[40] More explicitly testing for induced innovation is difficult because we do not observe what a firm does with its R&D expenditures. One way to proceed may be to add interactions of R&D expenditures and input prices to the laws of motion in equations (7) and (8). We leave this to future research.

Total technological change and its components. As productivity is multi-dimensional, we take total technological change to be $\epsilon_{Ljt-2}\Delta\omega_{Ljt} + \Delta\omega_{Hjt}$. Taken together, labor-augmenting and Hicks-neutral technological change cause output to grow by, on average, between 0.7% in industry 7 and 7.2% and 7.3% in industries 6 and 4, as can be seen in column (7) of Table 8. Across all industries, total technological change causes output to grow by 3.0% per year.

The output effect of labor-augmenting technological change $\epsilon_{Ljt-2}\Delta\omega_{Ljt}$ and Hicks-neutral technological change $\Delta\omega_{Hjt}$ are positively correlated in eight industries while the correlation is zero or slightly negative in two industries (column (8)). The correlation between labor-augmenting productivity in output terms $\epsilon_{Ljt-2}\omega_{Ljt}$ and Hicks-neutral productivity ω_{Hjt} is positive in all industries. Overall, our estimates not only provide evidence that productivity is multi- instead of single-dimensional but also suggest that the various components of productivity are intertwined.

9 An aggregate productivity growth decomposition

In quantifying labor-augmenting and Hicks-neutral technological change in Sections 6 and 8, we leverage our firm-level panel data to follow individual firms over time. In this section, we complement our findings by analyzing the aggregate productivity of the Spanish manufacturing sector and its growth over our sample period. To obtain insight into the drivers of growth, we decompose aggregate productivity growth along the lines of Olley & Pakes (1996).

Aggregate productivity $\phi_t = \sum_j \mu_{jt}\phi_{jt}$ in period t is a weighted average of the productivity of individual firms, where ϕ_{jt} is a measure of the productivity of firm j in period t and μ_{jt} is its weight. We separately examine labor-augmenting productivity in output terms $\epsilon_{Ljt-2}\omega_{Ljt}$, Hicks-neutral productivity ω_{Hjt}, and total productivity $\epsilon_{Ljt-2}\omega_{Ljt}+\omega_{Hjt}$. Throughout the weight $\mu_{jt} = (p_{jt}+y_{jt})/\sum_j(p_{jt}+y_{jt})$ is the share of the log of sales of firm j in period t.

Following Olley & Pakes (1996) and Melitz & Polanec (2015), we decompose the growth in aggregate productivity from period t_1 to period t_2 as

$$\begin{aligned}\Delta\phi = \phi_{t_2} - \phi_{t_1} &= \left(\phi^S_{t_2} - \phi^S_{t_1}\right) + \mu^E_{t_2}\left(\phi^E_{t_2} - \phi^S_{t_2}\right) + \mu^X_{t_1}\left(\phi^S_{t_1} - \phi^X_{t_1}\right) \\ &= \left(\bar{\phi}^S_{t_2} - \bar{\phi}^S_{t_1}\right) + N^S\left(Cov\left(\frac{\mu_{jt_2}}{\mu^S_{t_2}}, \phi_{jt_2}\right) - Cov\left(\frac{\mu_{jt_1}}{\mu^S_{t_1}}, \phi_{jt_1}\right)\right) + \mu^E_{t_2}\left(\phi^E_{t_2} - \phi^S_{t_2}\right) + \mu^X_{t_1}\left(\phi^S_{t_1} - \phi^X_{t_1}\right),\end{aligned} \tag{26}$$

where S, E, and X indexes the group of survivors, entrants, and exitors, respectively. $\mu^G_t = \sum_{j\in G}\mu_{jt}$ is the total weight of group G in period t, $\phi^G_t = \sum_{j\in G}\frac{\mu_{jt}}{\mu^G_t}\phi_{jt}$ is the weighted average restricted to group G, $\bar{\phi}^G_t = \frac{1}{N^G}\sum_{j\in G}\phi_{jt}$ is the unweighted average restricted to group G, and N^G is the number of firms in group G. In the first line of the decomposition, the first term captures the contribution of survivors to aggregate productivity growth, the second that of entrants, and the third that of exitors. The second line further decomposes the contribution of survivors to aggregate productivity growth into a shift in the distribution of productivity (first term) and a change in covariance that captures reallocation (second term).

As the decomposition pertains to the population of firms, applying it to the sample of firms in our firm-level panel data is subject to a caveat. As before we account for the survey design by replicating the subsample of small firms. We classify a firm as a survivor if it enters the industry in or before period t_1 and does not exit in or before period t_2. We further classify a firm as an entrant if it enters the industry after period t_1 and as an exitor if it exits the industry in or before period t_2. Due to attrition and the periodic addition of new firms to the sample, we observe productivity for a subset of survivors in period t_1 and for another subset of survivors in period t_2. Because we average over potentially quite different subsets of firms, especially if periods t_1 and t_2 are far apart, our estimates of the various terms in the decomposition may be noisy.

We report the change in aggregate productivity and its decomposition in equation (26) in Table 9 for the period 1992 to 2006 and the three subperiods 1992 to 1996, 1997 to 2001, and 2002 to 2006. The change in aggregate productivity in column (1) is consistent with our findings in Sections 6 and 8. Aggregate labor-augmenting productivity in output terms grew by 21.7% for the period 1992 to 2006 or about 1.6% per year. Aggregate Hicks-neutral productivity grew by 19.7% or about 1.4% per year and aggregate total productivity by 41.4% or about 3.0% per year. For the later subperiods, technological change appears to have slowed down, in particular in case of aggregate Hicks-neutral and total productivity.

Turning to the decomposition in columns (2), (5) and (6), survivors account for most of the change in aggregate total productivity and its components, again in line with our findings in Sections 6 and 8.[41] With the possible exception of entrants for the subperiod 1992 to 1996, the contribution of entrants and exitors appears to be limited. Honing in on survivors, shifts in the distribution of productivity are substantially more important than changes on the covariance, as can be from columns (3) and (4). The contribution of reallocation to the change in aggregate total productivity and its components is sometimes positive and sometimes negative and mostly small.

10 Capital-augmenting technological change

As discussed in Section 3, the evolution of the relative quantities and prices of the various factors of production provides no evidence for capital-augmenting technological change. Our leading specification therefore restricts the productivities of capital and materials to

[41] Survivors account for less of the change for the period 1992 to 2006 than for the three subperiods simply because the definition of survivor is more demanding if periods t_1 and t_2 are further apart.

change at the same rate and in lockstep with Hicks-neutral technological change. A more general specification allows for capital-augmenting productivity ω_{Kjt} so that the production function in equation (6) becomes

$$Y_{jt} = \left[\beta_K \left(\exp(\omega_{Kjt})K_{jt}\right)^{-\frac{1-\sigma}{\sigma}} + \left(\exp(\omega_{Ljt})L^*_{jt}\right)^{-\frac{1-\sigma}{\sigma}} + \beta_M \left(M^*_{jt}\right)^{-\frac{1-\sigma}{\sigma}}\right]^{-\frac{\nu\sigma}{1-\sigma}} \exp(\omega_{Hjt})\exp(e_{jt}). \tag{27}$$

We explore the role of capital-augmenting technological change in our data in two ways.

First, we follow Raval (2013) and parts of the previous literature on estimating aggregate production functions (see Antràs (2004) and the references therein) and assume that capital is a static input that is chosen each period to maximize short-run profits. In analogy to equation (15), we recover (conveniently rescaled) capital-augmenting productivity $\widetilde{\omega}_{Kjt} = (1-\sigma)\omega_{Kjt}$ as

$$\begin{aligned} \widetilde{\omega}_{Kjt} &= \widetilde{\gamma}_K + m_{jt} - k_{jt} + \sigma(p_{Mjt} - p_{Kjt}) + (1-\sigma)\gamma_1(S_{Ojt}) \\ &\equiv \widetilde{h}_K(m_{jt} - k_{jt}, p_{Mjt} - p_{Kjt}, S_{Ojt}), \end{aligned} \tag{28}$$

where $\widetilde{\gamma}_K = -\sigma \ln\left(\frac{\beta_M}{\beta_K}\right)$ and we use the user cost in our data as a rough measure of the price of capital p_{Kjt}. Using our leading estimates from Section 6, we recover the capital-augmenting productivity $\omega_{Kjt} = \frac{\widetilde{\omega}_{Kjt}}{1-\sigma}$ of firm j in period t up to an additive constant; in what follows, we use ω_{Kjt} to denote the de-meaned capital-augmenting productivity.[42] $\Delta\omega_{Kjt} \approx \frac{\exp(\omega_{Kjt})K_{jt-1} - \exp(\omega_{Kjt-1})K_{jt-1}}{\exp(\omega_{Kjt-1})K_{jt-1}}$ in column (1) of Table 10 approximates the rate of growth of a firm's effective capital stock $\exp(\omega_{Kjt-1})K_{jt-1}$ and $\epsilon_{Kjt-2}\Delta\omega_{Kjt}$ in column (2) the rate of growth of the firm's output Y_{jt-1}, where ϵ_{Kjt-2} is the elasticity of output with respect to the firm's effective capital stock (see Appendix D). As can be seen from column (1), capital-augmenting productivity grows slowly, on average, with rates of growth of 0.8% per year in industry 6, 2.2% in industry 10, and 5.6% in industry 1. The rate of growth is negative in the remaining seven industries. The growth of capital-augmenting productivity is especially underwhelming in comparison to the growth of labor-augmenting productivity (see again column (1) of Table 5). The output effect of capital-augmenting technological change in column (2) is also close to zero in all industries, although this likely reflects the fact that capital is not a static input. As the user cost excludes adjustment costs, it falls short of the shadow price of capital, and using it drives down the elasticity of output with respect to the firm's effective capital stock.

Second, we return to the usual setting in the literature following Olley & Pakes (1996)

[42] As an alternative to plugging our leading estimates from Section 6 into equation (28), in the Online Appendix we use equation (28) to form the analog to our first estimation equation (17):

$$\begin{gathered} m_{jt} - k_{jt} = -\sigma(p_{Mjt} - p_{Kjt}) - (1-\sigma)\gamma_1(S_{Ojt}) \\ + \widetilde{g}_{Kt-1}(\widetilde{h}_K(m_{jt-1} - k_{jt-1}, p_{Mjt-1} - p_{Kjt-1}, S_{Ojt-1}), R_{jt-1}) + \widetilde{\xi}_{Kjt}. \end{gathered}$$

Consistent with measurement error in p_{Kjt}, the resulting estimates of σ are very noisy and severely biased toward zero.

and allow the choice of capital to have dynamic implications. We follow parts of the previous literature on estimating aggregate production functions and proxy for ω_{Kjt} by a time trend $\delta_K t$. Our second estimation equation (20) remains unchanged except that

$$X_{jt} = \beta_K \left(\exp(\delta_K t) K_{jt}\right)^{-\frac{1-\sigma}{\sigma}} + \beta_M \left(M_{jt} \exp\left(\gamma_1(S_{Ojt})\right)\right)^{-\frac{1-\sigma}{\sigma}} \left(\frac{1 - S_{Mjt}}{S_{Mjt}} \lambda_1(S_{Tjt}) + 1\right).$$

Columns (3)–(7) of Table 10 summarize the resulting estimates of β_K, ν, and δ_K. The estimates of β_K and ν are very comparable to those in Table 5. Moreover, the insignificant time trend leaves little room for capital-augmenting technological change in our data.

In sum, in line with the patterns in the data described in Section 3, there is little, if any, evidence for capital-augmenting technological change in our data. Of course, our ways of exploring the role of capital-augmenting technological change are less than ideal in that they either rest on the assumption that capital is a static input or abstract from firm-level heterogeneity in capital-augmenting productivity. An important question is therefore whether our approach can be extended to treat capital-augmenting productivity on par with labor-augmenting and Hicks-neutral productivity.

Recovering a third component of productivity, at a bare minimum, requires a third decision besides labor and materials to invert. Investment is a natural candidate. In contrast to the demand for labor and materials, however, investment depends on the details of the firm's dynamic programming problem. There are two principal difficulties. First, one has to prove that the observed demands for labor and materials along with investment are jointly invertible for unobserved capital-augmenting, labor-augmenting, and Hicks-neutral productivity. Second, the inverse functions $\widetilde{h}_K(\cdot)$, $\widetilde{h}_L(\cdot)$, and $h_H(\cdot)$ are high-dimensional. Thus, estimating these functions nonparametrically is demanding on the data. In ongoing work, Zhang (2015) proposes combining a parametric inversion that exploits the parameter restrictions between production and input demand functions similar to our paper with a nonparametric inversion of investment similar to Olley & Pakes (1996).

11 Conclusions

Technological change can increase the productivity of capital, labor, and the other factors of production in equal terms, or it can be biased towards a specific factor. In this paper, we directly assess the bias of technological change by measuring, at the level of the individual firm, how much of technological change is labor augmenting and how much of it is Hicks neutral.

To this end, we develop a dynamic model of the firm in which productivity is multi-dimensional. At the center of the model is a CES production function that parsimoniously yet robustly relates the relative quantities of materials and labor to their relative prices and labor-augmenting productivity. To properly isolate and measure labor-augmenting productivity, we account for other factors that impact this relationship, in particular, outsourcing and adjustment costs on permanent labor.

We apply our estimator to an unbalanced panel of 2375 Spanish manufacturing firms in ten industries from 1990 to 2006. Our estimates indicate limited substitutability between the various factors of production. This calls into question whether the widely-used Cobb-Douglas production function with its unitary elasticity of substitution adequately represents firm-level production processes.

Our estimates provide clear evidence that technological change is biased. Ceteris paribus labor-augmenting technological change causes output to grow, on average, in the vicinity of 1.5% per year. While skill upgrading explains some of the growth of labor augmenting productivity, in many industries labor-augmenting productivity grows because workers with a given set of skills become more productive over time. In short, our estimates cast doubt on the assumption of Hicks-neutral technological change that underlies many of the standard techniques for measuring productivity and estimating production functions.

At the same time, however, our estimates do not validate the assumption that technological change is purely labor augmenting that plays a central role in the literature on economic growth. In addition to labor-augmenting technological change, our estimates show that Hicks-neutral technological change causes output to grow, on average, in the vicinity of 1.5% per year.

While we are primarily interested in measuring how much of technological change is labor augmenting and how much of it is Hicks neutral, we also use our estimates to illustrate the consequences of biased technological change beyond the growth of output. In particular, we show that it is the primary driver of the decline of the aggregate share of labor in the Spanish manufacturing sector over our sample period. An interesting avenue for future research is to investigate the implications of biased technological change for employment. Recent research points to biased technological change as a key driver of the diverging experiences of the continental European, U.S., and U.K. economies during the 1980s and 1990s (Blanchard 1997, Caballero & Hammour 1998, Bentolila & Saint-Paul 2004, McAdam & Willman 2013). Our estimates lend themselves to decomposing firm-level changes in employment into displacement, substitution, and output effects and to compare these effects between labor-augmenting and Hicks-neutral technological change. This may be helpful for better understanding and predicting the evolution of employment as well as for designing labor market and innovation policies in the presence of biased technological change.

Appendix A Proof of proposition 1

Rewriting the ratio of first-order conditions (3) yields

$$0 = \ln \frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial L_{jt}} + \omega_{Ljt} - \ln \frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial M_{jt}} + p_{Mjt} - w_{jt}$$
$$= f(m_{jt} - l_{jt}, p_{Mjt} - w_{jt}, \omega_{Ljt}).$$

Differentiating the so-defined function $f(\cdot)$ yields

$$\frac{\partial f(m_{jt} - l_{jt}, p_{Mjt} - w_{jt}, \omega_{Ljt})}{\partial(m_{jt} - l_{jt})}$$
$$= \left(- \frac{\frac{\partial^2 H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial L_{jt}^2}}{\frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial L_{jt}}} + \frac{\frac{\partial^2 H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial M_{jt} \partial L_{jt}}}{\frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial M_{jt}}} \right) \exp(\omega_{Ljt} - (m_{jt} - l_{jt}))$$
$$= \frac{H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right) \frac{\partial^2 H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial M_{jt} \partial L_{jt}}}{\frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial L_{jt}} \frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial M_{jt}}}$$
$$= \frac{1}{\sigma\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt}))\right)},$$

where the second equality uses that $H(\exp(\omega_{Ljt})L_{jt}, M_{jt})$ is homogeneous of degree one and the third equality uses that the elasticity of substitution between materials and labor (Chambers 1988, equation (1.13)) for the production function in equation (1) simplifies to

$$\sigma\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt}))\right) = \frac{\frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial L_{jt}} \frac{\partial H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial M_{jt}}}{H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right) \frac{\partial^2 H\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt})), 1\right)}{\partial M_{jt} \partial L_{jt}}}.$$

Similarly,

$$\frac{\partial f(m_{jt} - l_{jt}, p_{Mjt} - w_{jt}, \omega_{Ljt})}{\partial(p_{Mjt} - w_{jt})} = 1,$$
$$\frac{\partial f(m_{jt} - l_{jt}, p_{Mjt} - w_{jt}, \omega_{Ljt})}{\partial \omega_{Ljt}} = - \frac{1}{\sigma\left(\exp(\omega_{Ljt} - (m_{jt} - l_{jt}))\right)} + 1.$$

By the implicit function theorem, around a point $(m_{jt}^0 - l_{jt}^0, p_{Mjt}^0 - w_{jt}^0, \omega_{Ljt}^0)$ satisfying $f(m_{jt}^0 - l_{jt}^0, p_{Mjt}^0 - w_{jt}^0, \omega_{Ljt}^0) = 0$, there exists a continuously differentiable function $m_{jt} - l_{jt} = g(p_{Mjt} - w_{jt}, \omega_{Ljt})$ such that $f(g(p_{Mjt} - w_{jt}, \omega_{Ljt}), p_{Mjt} - w_{jt}, \omega_{Ljt}) = 0$ and

$$\frac{\partial g(p_{Mjt}^0 - w_{jt}^0, \omega_{Ljt}^0)}{\partial(p_{Mjt} - w_{jt})} = - \frac{\frac{\partial f(m_{jt}^0 - l_{jt}^0, p_{Mjt}^0 - w_{jt}^0, \omega_{Ljt}^0)}{\partial(p_{Mjt} - w_{jt})}}{\frac{\partial f(m_{jt}^0 - l_{jt}^0, p_{Mjt}^0 - w_{jt}^0, \omega_{Ljt}^0)}{\partial(m_{jt} - l_{jt})}} = -\sigma\left(\exp(\omega_{Ljt}^0 - (m_{jt}^0 - l_{jt}^0))\right),$$
$$\frac{\partial g(p_{Mjt}^0 - w_{jt}^0, \omega_{Ljt}^0)}{\partial \omega_{Ljt}} = - \frac{\frac{\partial f(m_{jt}^0 - l_{jt}^0, p_{Mjt}^0 - w_{jt}^0, \omega_{Ljt}^0)}{\partial \omega_{Ljt}}}{\frac{\partial f(m_{jt}^0 - l_{jt}^0, p_{Mjt}^0 - w_{jt}^0, \omega_{Ljt}^0)}{\partial(m_{jt} - l_{jt})}} = 1 - \sigma\left(\exp(\omega_{Ljt}^0 - (m_{jt}^0 - l_{jt}^0))\right).$$

The first-order Taylor series for $m_{jt} - l_{jt} = g(p_{Mjt} - w_{jt}, \omega_{Ljt})$ around the point $(m^0_{jt} - l^0_{jt}, p^0_{Mjt} - w^0_{jt}, \omega^0_{Ljt})$ follows immediately.

Appendix B Data

We observe firms for a maximum of 17 years between 1990 and 2006. We restrict the sample to firms with at least three years of data on all variables required for estimation. The number of firms with 3, 4,..., 17 years of data is 313, 240, 218, 215, 207, 171, 116, 189, 130, 89, 104, 57, 72, 94, and 160, respectively. Table A1 gives the industry definitions along with their equivalent definitions in terms of the ESEE, National Accounts, and ISIC classifications (columns (1)–(3)). Based on the National Accounts in 2000, we further report the shares of the various industries in the total value added of the manufacturing sector (column (4)).

In what follows we define the variables we use. We begin with the variables that are relevant for our main analysis.

- *Investment.* Value of current investments in equipment goods (excluding buildings, land, and financial assets) deflated by the price index of investment. The price of investment is the equipment goods component of the index of industry prices computed and published by the Spanish Ministry of Industry. By measuring investment in operative capital we avoid some of the more severe measurement issues of the other assets.

- *Capital.* Capital at current replacement values $\widetilde{K}_{jt}$ is computed recursively from an initial estimate and the data on current investments in equipment goods $\widetilde{I}_{jt}$. We update the value of the past stock of capital by means of the price index of investment P_{It} as $\widetilde{K}_{jt} = (1-\delta)\frac{P_{It}}{P_{It-1}}\widetilde{K}_{jt-1} + \widetilde{I}_{jt-1}$, where δ is an industry-specific estimate of the rate of depreciation. Capital in real terms is obtained by deflating capital at current replacement values by the price index of investment as $K_{jt} = \frac{\widetilde{K}_{jt}}{P_{It}}$.

- *Labor.* Total hours worked computed as the number of workers times the average hours per worker, where the latter is computed as normal hours plus average overtime minus average working time lost at the workplace.

- *Materials.* Value of intermediate goods consumption (including raw materials, components, energy, and services) deflated by a firm-specific price index of materials.

- *Output.* Value of produced goods and services computed as sales plus the variation of inventories deflated by a firm-specific price index of output.

- *Wage.* Hourly wage cost computed as total labor cost including social security payments divided by total hours worked.

- *Price of materials.* Firm-specific price index for intermediate consumption. Firms are asked about the price changes that occurred during the year for raw materials, components, energy, and services. The price index is computed as a Paasche-type index of the responses.

- *Price of output.* Firm-specific price index for output. Firms are asked about the price changes they made during the year in up to 5 separate markets in which they operate. The price index is computed as a Paasche-type index of the responses.

- *Demand shifter.* Firms are asked to assess the current and future situation of the main market in which they operate. The demand shifter codes the responses as 0, 0.5, and 1 for slump, stability, and expansion, respectively.

- *Share of temporary labor.* Fraction of workers with fixed-term contracts and no or small severance pay.

- *Share of outsourcing.* Fraction of customized parts and pieces that are manufactured by other firms in the value of the firm's intermediate goods purchases.

- *R&D expenditures.* R&D expenditures include the cost of intramural R&D activities, payments for outside R&D contracts with laboratories and research centers, and payments for imported technology in the form of patent licensing or technical assistance, with the various expenditures defined according to the OECD Oslo and Frascati manuals.

We next turn to additional variables that we use for descriptive purposes, extensions, and robustness checks.

- *User cost of capital.* Computed as $P_{It}(r_{jt} + \delta - CPI_t)$, where P_{It} is the price index of investment, r_{jt} is a firm-specific interest rate, δ is an industry-specific estimate of the rate of depreciation, and CPI_t is the rate of inflation as measured by the consumer price index.

- *Skill mix.* Fraction of non-production employees (white collar workers), workers with an engineering degree (engineers), and workers with an intermediate degree (technicians).

- *Region.* Dummy variables corresponding to the 19 Spanish autonomous communities and cities where employment is located if it is located in a unique region and another dummy variable indicating that employment is spread over several regions.

- *Product submarket.* Dummy variables corresponding to a finer breakdown of the 10 industries into subindustries (restricted to subindustries with at least 5 firms, see column (5) of Table A1).

- *Technological sophistication.* Dummy variable that takes the value one if the firm uses digitally controlled machines, robots, CAD/CAM, or some combination of these procedures.

- *Identification between ownership and control.* Dummy variable that takes the value one if the owner of the firm or the family of the owner hold management positions.

- *Age.* Years elapsed since the foundation of the firm with a maximum of 40 years.

- *Firm size.* Number of workers in the year the firm enters the sample.

Appendix C Estimation

Unknown functions. The functions $\widetilde{g}_{L1}(\widetilde{h}_L(\cdot))$, $\widetilde{g}_{L2}(\widetilde{h}_L(\cdot), r_{jt-1})$, $g_{H1}(h_H(\cdot))$, and $g_{H2}(h_H(\cdot), r_{jt-1})$ that are part of the conditional expectation functions $\widetilde{g}_{Lt-1}(\widetilde{h}_L(\cdot), R_{jt-1})$ and $g_{Ht-1}(h_H(\cdot), R_{jt-1})$ are unknown and must be estimated nonparametrically, as must be the absolute value of the price elasticity $\eta(p_{jt}, D_{jt})$ and the correction terms $\lambda_1(S_{Tjt})$, $\lambda_2(S_{Tjt})$, and $\gamma_1(S_{Ojt})$. Following Wooldridge (2004), we model an unknown function $q(v)$ of one variable v by a univariate polynomial of degree Q. We model an unknown function $q(u, v)$ of two variables u and v by a complete set of polynomials of degree Q (see Judd 1998). Unless otherwise noted, we omit the constant in $q(\cdot)$ and set $Q = 3$ in the remainder of this paper.

Starting with the conditional expectation functions, we specify $\widetilde{g}_{L1}(\widetilde{h}_L(\cdot)) = q(\widetilde{h}_L(\cdot) - \widetilde{\gamma}_L)$, $\widetilde{g}_{L2}(\widetilde{h}_L(\cdot), r_{jt}) = q_0 + q(\widetilde{h}_L(\cdot) - \widetilde{\gamma}_L, r_{jt})$, $g_{H1}(h_H(\cdot)) = q(h_H(\cdot) - \gamma_H)$, and $g_{H2}(h_H(\cdot), r_{jt}) = q_0 + q(h_H(\cdot) - \gamma_H, r_{jt})$, where q_0 is a constant and the function $q(\cdot)$ is modeled as described above. Without loss of generality, we absorb $\widetilde{\gamma}_L$ and γ_H into the overall constants of our estimation equations. Turning to the absolute value of the price elasticity, to impose the theoretical restriction $\eta(p_{jt}, D_{jt}) > 1$, we specify $\eta(p_{jt}, D_{jt}) = 1 + \exp(q(p_{jt}, D_{jt}))$, where the function $q(\cdot)$ is modeled as described above except that we suppress terms involving D_{jt}^2 and D_{jt}^3. Turning to the correction terms, we specify $\lambda_1(S_{Tjt}) = q(\ln S_{Tjt})$ and $\lambda_2(S_{Tjt}) = q(\ln S_{Tjt})$ in industries 2, 3, and 10 and $\lambda_1(S_{Tjt}) = q(\ln(1 - S_{Tjt}))$ and $\lambda_2(S_{Tjt}) = q(\ln(1 - S_{Tjt}))$ in the remaining industries.[43] Finally, we specify $\gamma_1(S_{Ojt}) = q(S_{Ojt})$; this ensures that $\gamma_1(S_{Ojt}) = 0$ if $S_{Ojt} = 0$ in line with the normalization $\Gamma(M_{Ijt}, 0) = M_{Ijt}$.

Parameters and instruments. Our first estimation equation (17) has 36 parameters: constant, σ, 15 parameters in $\widetilde{g}_{L0}(t-1)$ (time dummies), 3 parameters in $\widetilde{g}_{L1}(\widetilde{h}_L(\cdot))$, 10 parameters in $\widetilde{g}_{L2}(\widetilde{h}_L(\cdot), r_{jt-1})$, 3 parameters in $\lambda_2(S_{Tjt})$, and 3 parameters in $\gamma_1(S_{Ojt})$.

Our instrumenting strategy is adapted from Doraszelski & Jaumandreu (2013) and we refer the reader to Doraszelski & Jaumandreu (2013) and the references therein for a discussion of the use of polynomials for instruments. We use the constant, 15 time dummies, the dummy for performers $1(R_{jt-1} > 0)$, the demand shifter D_{jt}, and a univariate polynomial in $\ln S_{Ojt-1} + m_{jt-1}$ interacted with $1(S_{Ojt-1} > 0)$ (3 instruments). We further use a complete set of polynomials in l_{jt-1}, m_{jt-1}, and $p_{Mjt-1} - w_{jt-1}$ interacted with the dummy for nonperformers $1(R_{jt-1} = 0)$ (19 instruments). In industries 5 and 8 we replace $p_{Mjt-1} - w_{jt-1}$ by p_{Mjt-1} in the complete set of polynomials. Finally, we use a complete set of polynomials in l_{jt-1}, m_{jt-1}, and $p_{Mjt-1} - w_{jt-1}$ and r_{jt-1} interacted with the dummy for performers $1(R_{jt-1} > 0)$ (34 instruments). This yields a total of 74 instruments and $74 - 36 = 38$ degrees of freedom (see column (4) of Table 4).

After imposing the estimated values from equation (17), our second estimation equation (20) has 40 parameters: constant, β_K, ν, 15 parameters in $g_{H0}(t-1)$ (time dummies), 3 parameters in $g_{H1}(h_H(\cdot))$, 10 parameters in $g_{H2}(h_H(\cdot), r_{jt-1})$, 3 parameters in $\lambda_1(S_{Tjt})$, and 6 parameters in $\eta(p_{jt}, D_{jt})$.

As before, we use polynomials for instruments. We use the constant, 15 time dummies, the dummy for performers $1(R_{jt-1} > 0)$, the demand shifter D_{jt}, a univariate polynomial in p_{jt-1} (3 instruments), a univariate polynomial in $p_{Mjt-1} - p_{jt-1}$ (3 instruments), and a univariate polynomial in k_{jt} (3 instruments). We also use a complete set of polynomials in $M_{jt-1}\frac{1-S_{Mjt-1}}{S_{Mjt-1}}$ and K_{jt-1} interacted with the dummy for nonperformers $1(R_{jt-1} = 0)$ (9 instruments). Finally, we use a complete set of polynomials in $M_{jt-1}\frac{1-S_{Mjt-1}}{S_{Mjt-1}}$ and K_{jt-1} (9 instruments) and a univariate polynomial in r_{jt-1} interacted with the dummy for performers $1(R_{jt-1} > 0)$ (3 instruments). This yields a total of 48 instruments and $48 - 40 = 8$ degrees

[43] To incorporate skill upgrading, we instead specify $\lambda_1(S_{Tjt}, \Theta_{jt}) = q(\ln S_{Tjt}, \ln \Theta_{jt})$ and $\lambda_2(S_{Tjt}, \Theta_{jt}) = q(\ln S_{Tjt}, \ln \Theta_{jt})$ in industries 2, 3, and 10 and $\lambda_1(S_{Tjt}, \Theta_{jt}) = q(\ln(1 - S_{Tjt}), \ln \Theta_{jt})$ and $\lambda_2(S_{Tjt}, \Theta_{jt}) = q(\ln(1 - S_{Tjt}), \ln \Theta_{jt})$ in the remaining industries, where the function $q(\cdot)$ is modeled as described above except that we suppress terms involving $(\ln \Theta_{jt})^2$ and $(\ln \Theta_{jt})^3$.

of freedom in industries 1, 2, 3, 6, 7, 9, and 10 (see column (3) of Table 7). In industries 4, 5, and 8, we add a univariate polynomial in $\ln(1 - S_{Tjt-1})$ (3 instruments). We replace the univariate polynomial in k_{jt} by k_{jt} in industries 4 and 8 and we drop D_{jt} in industry 5.

Estimation. From the GMM problem in equation (21) with weighting matrix $\widehat{W}_L = \left[\frac{1}{N}\sum_j A_{Lj}(z_j)A_{Lj}(z_j)'\right]^{-1}$ we first obtain a consistent estimate $\widehat{\theta}_L$ of θ_L. This first step is the NL2SLS estimator of Amemiya (1974). In the second step, we compute the optimal estimate with weighting matrix $\widehat{W}_L = \left[\frac{1}{N}\sum_j A_{Lj}(z_j)\nu_{Lj}(\widehat{\theta}_L)\nu_{Lj}(\widehat{\theta}_L)'A_{Lj}(z_j)'\right]^{-1}$. Throughout the paper, we report standard errors that are robust to heteroskedasticity and autocorrelation.

Implementation. Code for our estimator is available from the authors upon request along with instructions for obtaining the data. We use Gauss 14.0.9 and Optmum 3.1.7.

To the reduce the number of parameters to search over in the GMM problem in equation (21), we "concentrate out" the parameters that enter it linearly (Wooldridge 2010, p. 435). To guard against local minima, we have extensively searched over the remaining parameters, often using preliminary estimates to narrow down the range of these parameters.

Testing. The value of the GMM objective function for the optimal estimator, multiplied by N, has a limiting χ^2 distribution with $Q - P$ degrees of freedom, where Q is the number of instruments and P the number of parameters to be estimated. We use it as a test for overidentifying restrictions or validity of the moment conditions.

Appendix D Output effect

Direct calculation starting from equation (6) yields the elasticity of output with respect to a firm's effective labor force:

$$\begin{aligned}\epsilon_{Ljt} &= \frac{\partial Y_{jt}}{\partial \exp(\omega_{Ljt})L^*_{jt}}\frac{\exp(\omega_{Ljt})L^*_{jt}}{Y_{jt}} \\ &= \frac{\nu\left(\exp(\omega_{Ljt})L^*_{jt}\right)^{-\frac{1-\sigma}{\sigma}}}{\beta_K K_{jt}^{-\frac{1-\sigma}{\sigma}} + \left(\exp(\omega_{Ljt})L^*_{jt}\right)^{-\frac{1-\sigma}{\sigma}} + \beta_M\left(M^*_{jt}\right)^{-\frac{1-\sigma}{\sigma}}}. \end{aligned} \tag{29}$$

Using equation (15) to substitute for ω_{Ljt} and simplifying we obtain

$$\epsilon_{Ljt} = \frac{\nu\frac{1-S_{Mjt}}{S_{Mjt}}\lambda_1(S_{Tjt})}{\frac{\beta_K}{\beta_M}\left(\frac{K_{jt}}{M_{jt}\exp\left(\gamma_1(S_{Ojt})\right)}\right)^{-\frac{1-\sigma}{\sigma}} + \frac{1-S_{Mjt}}{S_{Mjt}}\lambda_1(S_{Tjt}) + 1}. \tag{30}$$

Recall from equation (12) that $\lambda_1(S_{Tjt}) = 1 + \frac{\Delta_{jt}}{1+\frac{W_{Tjt}}{W_{Pjt}}\frac{S_{Tjt}}{1-S_{Tjt}}}$, where Δ_{jt} is the gap between the wage of permanent workers W_{Pjt} and the shadow wage. To facilitate evaluating equation (30), we abstract from adjustment costs and set $\lambda_1(S_{Tjt}) = 1$.

Direct calculation starting from equation (27) also yields the elasticity of output with respect to a firm's effective capital stock:

$$\epsilon_{Kjt} = \frac{\partial Y_{jt}}{\partial \exp(\omega_{Kjt})K_{jt}} \frac{\exp(\omega_{Kjt})K_{jt}}{Y_{jt}}$$

$$= \frac{\nu \left(\exp(\omega_{Kjt})K_{jt}\right)^{-\frac{1-\sigma}{\sigma}}}{\left(\exp(\omega_{Kjt})K_{jt}\right)^{-\frac{1-\sigma}{\sigma}} + \left(\exp(\omega_{Ljt})L^*_{jt}\right)^{-\frac{1-\sigma}{\sigma}} + \beta_M \left(M^*_{jt}\right)^{-\frac{1-\sigma}{\sigma}}}$$

$$= \frac{\nu}{1 + \frac{P_{Mjt}M_{jt}}{P_{Kjt}K_{jt}} \left(\frac{1-S_{Mjt}}{S_{Mjt}} \lambda_1(S_{Tjt}) + 1\right)}, \tag{31}$$

where we use equations (15) and (28) to substitute for ω_{Ljt} and ω_{Kjt}, respectively. As with equation (30), we set $\lambda_1(S_{Tjt}) = 1$ to evaluate equation (31).

Appendix E Wage regression

As column (1) of Table A2 shows, the coefficient of variation for the (level of the) wage W_{jt} ranges from 0.35 to 0.50 across industries.[44] The variance decomposition in columns (2)–(4) shows that around one quarter of the overall variation is within firms across periods. The larger part of this variation is across firms.

To explore the source of this variation, we regress the (log of the) wage w_{jt} on the skill mix of a firm's labor force as given by the share of temporary (as opposed to permanent) labor, the share of white (as opposed to blue) collar workers, and the shares of engineers and technicians (as opposed to unskilled workers), time dummies, region dummies, product submarket dummies, the demand shifter, and an array of other firm characteristics, namely dummies for technological sophistication and identification of ownership and control as well as univariate polynomials of degree 3 in age and firm size.

To motivate this regression, assume that there are Q types of labor with wages W_{1jt}, W_{2jt}, ..., W_{Qjt} and write the wage as

$$W_{jt} = \sum_{q=1}^{Q} W_{qjt} S_{qjt} = W_{1jt} \left(1 + \sum_{q=2}^{Q} \left(\frac{W_{qjt}}{W_{1jt}} - 1\right) S_{qjt}\right),$$

where S_{qjt} is the share of labor of type q and $\sum_{q=1}^{Q} S_{qjt} = 1$. Because

$$w_{jt} \approx w_{1jt} + \sum_{q=2}^{Q} \left(\frac{W_{qjt}}{W_{1jt}} - 1\right) S_{qjt},$$

the coefficient on S_{qjt} in the wage regression is an estimate of the wage premium $\left(\frac{W_{qjt}}{W_{1jt}} - 1\right)$ of labor of type q over type 1. Because we do not have the joint distribution of skills (e.g., temporary white collar technician) in our data, we approximate it by the marginal distributions (e.g., share of temporary labor) and ignore higher-order terms. As columns (5)–(8) of Table A2 show, the estimated coefficients on the skill mix of a firm's labor force are often significant, have the expected signs, and are quite similar across industries. On

[44] The coefficient of variation for the price of materials ranges from 0.12 to 0.19 across industries.

average across industries, temporary workers earn 36% less than permanent workers, white collar workers earn 26% more than blue collar workers, engineers earn 85% more than unskilled workers, and technicians earn 23% more than unskilled workers.

The wage regression also shows that some, but by no means all variation in the wage is due to worker quality. To isolate the part of the wage that depends on the skill mix of a firm's labor force, we decompose the predicted wage $\widehat{w}_{jt}$ into a prediction $\widehat{w}_{Qjt}$ based on the skill mix and a prediction $\widehat{w}_{Cjt}$ based on the remaining variables. $\widehat{w}_{Qjt}$ and $\widehat{w}_{Cjt}$ are positively correlated. According to $R^2 = \frac{Var(\widehat{w}_{jt})}{Var(w_{jt})}$ in column (9), depending on the industry, the wage regression explains between 63% and 76% of the variation in the wage, with an average of 70%. The skill mix by itself explains between 2% and 20% of the variation in the wage, with an average of 10% (see $R_Q^2 = \frac{Var(\widehat{w}_{Qjt})}{Var(w_{jt})}$ in column (10)). In contrast, the remaining variables explain between 36% and 64% of the variation in the wage, with an average of 48% (see $R_C^2 = \frac{Var(\widehat{w}_{Cjt})}{Var(w_{jt})}$ in column (11)). The larger part of the variation in the wage therefore appears to be due to temporal and geographic differences in the supply of labor, the fact that firms operate in different product submarkets, and other firm characteristics.

In developing the quality index Θ_{jt}, we assume that there are Q types of permanent labor. We approximate the wage premium $\left(\frac{W_{Pqjt}}{W_{P1jt}} - 1\right)$ of permanent labor of type q over type 1 by the estimated coefficient on S_{qjt} in the wage regression and the share $S_{Pqjt} = \frac{L_{Pqjt}}{L_{Pjt}} = \frac{L_{Pqjt}}{L_{jt}} \Big/ \frac{L_{Pjt}}{L_{jt}}$ of permanent labor of type q by $\frac{S_{qjt}}{1-S_{Tjt}}$.

Appendix F Second-order approximation

Let $\Upsilon_{jt} = -\widetilde{\gamma}_L + (1-\sigma)(p_{Mjt} - w_{jt} + \omega_{Ljt}) + \sigma\lambda_2(S_{Tjt}) - (1-\sigma)\gamma_1(S_{Ojt})$ and $\Delta\Upsilon_{jt} = \Upsilon_{jt} - \Upsilon_{jt-1}$. Using equation (25) we write

$$\begin{aligned} S_{Ljt} - S_{Ljt-1} &= -S_{Ljt}(1 - S_{Ljt-1})\left(\exp\left(\Delta\Upsilon_{jt}\right) - 1\right) \\ &\approx -S_{Ljt}(1 - S_{Ljt-1})\left(\Delta\Upsilon_{jt} + \frac{1}{2}\left(\Delta\Upsilon_{jt}\right)^2\right), \end{aligned}$$

where we replace $\exp(\Delta\Upsilon_{jt}) - 1$ by its second-order Taylor series approximation around $\Delta\Upsilon_{jt} = 0$. We allocate the interactions in $(\Delta\Upsilon_{jt})^2$ in equal parts to the variables involved.

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Table 1: Descriptive statistics.

Industry	Obs.[a]	Firms[a]	Rates of growth[b] Output (s. d.)	Capital (s. d.)	Labor (s. d.)	Materials (s. d.)	Price (s. d.)	$\frac{M}{L}$ (s. d.)	$\frac{P_M}{W}$ (s. d.)	$\frac{M}{K}$ (s. d.)	$\frac{P_M}{P_K}$ (s. d.)
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
1. Metals and metal products	2365	313	0.045	0.051	0.008	0.030	0.017	0.022	-0.008	-0.021	0.049
			(0.235)	(0.192)	(0.161)	(0.327)	(0.052)	(0.316)	(0.176)	(0.373)	(0.099)
2. Non-metallic minerals	1270	163	0.046	0.057	0.010	0.041	0.012	0.031	-0.012	-0.016	0.043
			(0.228)	(0.212)	(0.177)	(0.285)	(0.058)	(0.272)	(0.147)	(0.333)	(0.104)
3. Chemical products	2168	299	0.060	0.062	0.015	0.044	0.008	0.029	-0.015	-0.019	0.044
			(0.228)	(0.182)	(0.170)	(0.274)	(0.055)	(0.250)	(0.153)	(0.313)	(0.141)
4. Agric. and ind. machinery	1411	178	0.031	0.040	-0.003	0.018	0.015	0.022	-0.015	-0.021	0.041
			(0.252)	(0.190)	(0.169)	(0.347)	(0.026)	(0.335)	(0.155)	(0.390)	(0.099)
5. Electrical goods	1505	209	0.059	0.041	0.010	0.048	0.008	0.038	-0.021	0.007	0.045
			(0.268)	(0.173)	(0.205)	(0.359)	(0.046)	(0.344)	(0.174)	(0.394)	(0.095)
6. Transport equipment	1206	161	0.060	0.043	0.004	0.051	0.008	0.047	-0.019	0.008	0.033
			(0.287)	(0.164)	(0.201)	(0.375)	(0.031)	(0.343)	(0.171)	(0.396)	(0.093)
7. Food, drink and tobacco	2455	327	0.023	0.047	0.003	0.012	0.021	0.009	-0.018	-0.035	0.049
			(0.206)	(0.177)	(0.169)	(0.286)	(0.054)	(0.295)	(0.176)	(0.328)	(0.116)
8. Textile, leather and shoes	2368	335	0.004	0.031	-0.015	-0.009	0.015	0.006	-0.021	-0.040	0.040
			(0.229)	(0.189)	(0.180)	(0.348)	(0.042)	(0.355)	(0.183)	(0.385)	(0.099)
9. Timber and furniture	1445	207	0.025	0.045	0.013	0.014	0.020	0.001	-0.019	-0.031	0.067
			(0.225)	(0.168)	(0.184)	(0.335)	(0.031)	(0.329)	(0.171)	(0.371)	(0.123)
10. Paper and printing products	1414	183	0.031	0.052	-0.001	0.013	0.017	0.014	-0.017	-0.039	0.046
			(0.187)	(0.221)	(0.149)	(0.252)	(0.074)	(0.247)	(0.159)	(0.326)	(0.122)

[a] Including $S_{Tjt} = L_{Tjt} = 0$.
[b] Computed for 1991 to 2006.

Table 2: Descriptive statistics (cont'd).

Industry	Temp. labor Obs. (%)	Temp. labor Share (s. d.)	Intrafirm max-min Share of temp. (s. d.)	Intrafirm max-min Hours worked[a] (s. d.)	Intrafirm max-min Hours per worker[a] (s. d.)	Outsourcing Obs. (%)	Outsourcing Share (s. d.)	With R&D Stable (%)	With R&D Occas. (%)	With R&D R&D intens. (s. d.)
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
1. Metal and metal products	1877	0.260	0.243	0.448	0.069	1014	0.200	56	109	0.012
	(79.4)	(0.221)	(0.197)	(0.360)	(0.090)	(42.9)	(0.193)	(17.9)	(34.8)	(0.018)
2. Non-metallic minerals	1018	0.231	0.232	0.482	0.065	316	0.177	20	62	0.011
	(80.2)	(0.207)	(0.183)	(0.403)	(0.063)	(24.9)	(0.179)	(12.3)	(38.0)	(0.022)
3. Chemical products	1722	0.170	0.203	0.446	0.043	924	0.146	121	85	0.026
	(79.4)	(0.176)	(0.185)	(0.427)	(0.038)	(42.6)	(0.183)	(40.5)	(28.4)	(0.034)
4. Agric. and ind. machinery	1069	0.189	0.227	0.485	0.086	808	0.288	64	62	0.022
	(75.8)	(0.181)	(0.181)	(0.419)	(0.166)	(57.3)	(0.263)	(36.0)	(34.8)	(0.026)
5. Electrical goods	1221	0.245	0.280	0.559	0.063	763	0.181	83	61	0.029
	(81.1)	(0.206)	(0.216)	(0.452)	(0.077)	(50.7)	(0.194)	(39.7)	(29.2)	(0.040)
6. Transport equipment	962	0.206	0.239	0.555	0.131	637	0.233	60	56	0.028
	(79.8)	(0.198)	(0.184)	(0.415)	(0.237)	(52.8)	(0.261)	(37.3)	(34.8)	(0.049)
7. Food, drink and tobacco	2067	0.276	0.266	0.468	0.058	514	0.142	65	86	0.007
	(84.2)	(0.237)	(0.215)	(0.343)	(0.065)	(20.9)	(0.172)	(19.9)	(26.3)	(0.022)
8. Textile, leather and shoes	1726	0.238	0.291	0.489	0.062	1214	0.252	44	85	0.017
	(79.2)	(0.260)	(0.244)	(0.402)	(0.086)	(51.3)	(0.237)	(13.1)	(25.4)	(0.031)
9. Timber and furniture	1175	0.320	0.326	0.523	0.056	535	0.183	21	44	0.010
	(81.3)	(0.226)	(0.234)	(0.387)	(0.076)	(37.0)	(0.201)	(10.1)	(21.3)	(0.017)
10. Paper and printing products	1024	0.155	0.221	0.425	0.057	679	0.273	17	48	0.015
	(72.4)	(0.145)	(0.196)	(0.346)	(0.065)	(48.0)	(0.253)	(9.3)	(26.2)	(0.028)

[a] Computed as difference in logs.

Table 3: Elasticity of substitution.

Industry	Obs.[a]	Firms[a]	OLS σ (s. e.)	OLS $\tilde{\delta}_L$ (s. e.)	GMM incl. m_{jt-1} as instr. σ (s. e.)	GMM incl. m_{jt-1} as instr. χ^2 (df)	GMM incl. m_{jt-1} as instr. p val.	GMM excl. m_{jt-1} as instr. σ (s. e.)	GMM excl. m_{jt-1} as instr. χ^2 (df)	GMM excl. m_{jt-1} as instr. p val.
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
1. Metals and metal products	2365	313	1.163	0.023	0.451	57.846	0.034	0.694	13.683	0.550
			(0.104)	(0.007)	(0.096)	(40)		(0.113)	(15)	
2. Non-metallic minerals	1270	163	1.227	0.038	0.643	46.068	0.234	0.603	11.299	0.731
			(0.119)	(0.008)	(0.086)	(40)		(0.126)	(15)	
3. Chemical products	2168	299	1.132	0.016	0.481	65.068	0.007	0.618	7.582	0.939
			(0.095)	(0.007)	(0.099)	(40)		(0.124)	(15)	
4. Agric. and ind. machinery	1411	178	1.239	0.019	0.502	56.166	0.046	0.598	8.500	0.902
			(0.166)	(0.008)	(0.114)	(40)		(0.103)	(15)	
5. Electrical goods	1505	209	1.402	0.017	0.469	60.674	0.019	0.458	17.457	0.292
			(0.163)	(0.009)	(0.108)	(40)		(0.108)	(15)	
6. Transport equipment	1206	161	1.161	0.029	1.204	48.449	0.169	0.512	7.740	0.934
			(0.218)	(0.011)	(0.089)	(40)		(0.162)	(15)	
7. Food, drink and tobacco	2455	327	1.421	0.015	0.614	70.492	0.002	0.707	15.088	0.445
			(0.094)	(0.008)	(0.063)	(40)		(0.084)	(15)	
8. Textile, leather and shoes	2368	335	1.846	0.001	0.059	55.178	0.056	0.724	18.453	0.240
			(0.169)	(0.100)	(0.077)	(40)		(0.162)	(15)	
9. Timber and furniture	1445	207	0.793	0.014	0.461	37.357	0.590	0.486	5.805	0.983
			(0.117)	(0.008)	(0.089)	(40)		(0.102)	(15)	
10. Paper and printing products	1414	183	1.120	0.026	0.609	51.798	0.100	0.854	7.300	0.949
			(0.107)	(0.008)	(0.057)	(40)		(0.077)	(15)	

[a] Including $S_{Tjt} = L_{Tjt} = 0$.

Table 4: Elasticity of substitution (cont'd).

			GMM			Sargan difference tests: w_{-1}		Sargan difference tests: $p_{M,-1}$		GMM with quality-corrected wage as instr.		
Industry	Obs.[a]	Firms[a]	σ (s. e.)	χ^2 (df)	p val.	χ^2 (df)	p val.	χ^2 (df)	p val.	σ (s. e.)	χ^2 (df)	p val.
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
1. Metals and metal products	1759	278	0.535	48.882	0.111	40.773	0.024	34.044	0.107	0.456	52.058	0.064
			(0.114)	(38)		(25)		(25)		(0.112)	(38)	
2. Non-metallic minerals	959	146	0.730	46.890	0.153	36.034	0.071	35.743	0.076	0.833	45.105	0.199
			(0.098)	(38)		(25)		(25)		(0.096)	(38)	
3. Chemical products	1610	269	0.696	46.154	0.171	33.225	0.126	32.183	0.153	0.695	48.889	0.111
			(0.102)	(38)		(25)		(25)		(0.072)	(38)	
4. Agric. and ind. machinery	979	164	0.607	42.420	0.286	29.398	0.248	25.684	0.425	0.762	44.227	0.225
			(0.196)	(38)		(25)		(25)		(0.206)	(38)	
5. Electrical goods	1147	191	0.592	46.782	0.155	38.951	0.037	32.376	0.147	0.624	44.592	0.214
			(0.123)	(38)		(25)		(25)		(0.125)	(38)	
6. Transport equipment	896	146	0.798	45.740	0.182	19.053	0.795	9.901	0.997	0.602	41.214	0.332
			(0.088)	(38)		(25)		(25)		(0.097)	(38)	
7. Food, drink and tobacco	1963	306	0.616	53.931	0.045	53.454	0.001	28.523	0.284	0.766	38.379	0.452
			(0.081)	(38)		(25)		(25)		(0.079)	(38)	
8. Textile, leather and shoes	1593	282	0.440	52.496	0.059	23.355	0.557	31.763	0.165	0.462	55.996	0.030
			(0.186)	(38)		(25)		(25)		(0.149)	(38)	
9. Timber and furniture	1114	188	0.438	39.204	0.416	28.979	0.265	22.059	0.632	0.497	36.687	0.530
			(0.093)	(38)		(25)		(25)		(0.094)	(38)	
10. Paper and printing products	938	162	0.525	44.508	0.217	23.642	0.540	19.822	0.756	0.449	43.009	0.265
			(0.088)	(38)		(25)		(25)		(0.085)	(38)	

[a] Excluding $S_{Tjt} = L_{Tjt} = 0$.

Table 5: Labor-augmenting technological change.

					Firms' R&D activities			Skill upgrading			
			$\epsilon_{L,-2}\omega_L$[a]		$\epsilon_{L,-2}\omega_L$	$\epsilon_{L,-2}\Delta\omega_L$		σ	χ^2 (df)	p val.	$\Delta\omega_L$
Industry	$\Delta\omega_L$	$\epsilon_{L,-2}\Delta\omega_L$	IQR	AC	R&D–No R&D	R&D	No R&D	(s. e.)			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
1. Metals and metal products	0.091	0.021	0.368	0.707	0.189	0.024	0.018	0.582	44.868	0.206	0.104
								(0.117)	(38)		
2. Non-metallic minerals	0.146	0.031	0.595	0.859	0.301	0.022	0.029	0.737	35.898	0.567	0.087
								(0.092)	(38)		
3. Chemical products	0.061	0.016	0.465	0.892	0.196	0.020	0.001	0.618	47.832	0.132	0.053
								(0.110)	(38)		
4. Agric. and ind. machinery	0.125	0.032	0.574	0.799	0.383	0.028	0.046	0.177	38.413	0.451	0.060
								(0.172)	(38)		
5. Electrical goods	0.216	0.021	0.526	0.854	0.298	0.022	0.011	0.488	48.365	0.121	0.179
								(0.129)	(38)		
6. Transport equipment	0.145	0.030	0.699	0.841	0.378	0.038	0.012	0.781	45.457	0.189	0.098
								(0.101)	(38)		
7. Food, drink and tobacco	0.017	0.006	0.455	0.874	0.063	0.009	0.006	0.655	53.981	0.045	-0.007
								(0.084)	(38)		
8. Textile, leather and shoes	0.038	0.009	0.389	0.844	0.186	0.012	0.009	0.120	41.931	0.304	0.000
								(0.168)	(38)		
9. Timber and furniture	0.067	0.002	0.241	0.613	0.100	0.007	0.001	0.528	37.674	0.484	-0.023
								(0.090)	(38)		
10. Paper and printing products	0.009	0.012	0.279	0.818	0.133	0.006	0.019	0.396	37.418	0.496	-0.011
								(0.082)	(38)		
All industries	0.100	0.016									0.042

[a] Without replication and weighting.

Table 6: Aggregate share of labor in variable cost.

Industry	Growth of labor share[a]	Decomposition[a]					
		$p_M - w$	ω_L	Temp. labor	Outsourcing	Reallocation	Residual
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
1. Metal and metal products	-0.107	0.011	-0.086	-0.034	0.011	-0.004	-0.004
2. Non-metallic minerals	-0.153	0.011	-0.130	-0.004	-0.002	-0.026	-0.003
3. Chemical products	-0.087	0.018	-0.152	0.041	-0.002	0.008	0.000
4. Agric. and ind. machinery	-0.046	0.031	-0.064	-0.026	-0.007	0.016	0.004
5. Electrical goods	-0.188	0.032	-0.178	0.021	-0.017	-0.040	-0.006
6. transport equipment	-0.066	0.024	-0.119	0.027	0.003	0.003	-0.003
7. Food, drink and tobacco	-0.060	0.022	-0.096	-0.013	0.024	-0.001	0.005
8. Textile, leather and shoes	-0.057	0.017	-0.051	-0.034	0.006	0.016	-0.011
9. Timber and furniture	-0.009	0.056	-0.046	-0.002	-0.002	-0.018	0.002
10. Paper and printing products	-0.059	0.021	-0.092	-0.001	-0.001	0.013	0.000

[a] Computed for 1991 to 2006.

Table 7: Distributional parameters, elasticity of scale, and price elasticity

Industry	GMM β_K (s. e.)	ν (s. e.)	χ^2 (df)	p val.	$\eta(p_{-1}, D_{-1})^a$	Lagrange-multiplier test $\chi^2(1)$	p val.
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
1. Metals and metal products	0.232	0.941	3.207	0.921	2.371	1.023	0.312
	(0.073)	(0.029)	(8)				
2. Non-metallic minerals	0.225	0.911	4.528	0.807	9.114	0.489	0.485
	(0.133)	(0.063)	(8)				
3. Chemical products	0.136	0.934	1.109	0.997	2.431	0.342	0.559
	(0.059)	(0.041)	(8)				
4. Agric. and ind. machinery	0.139	0.806	8.251	0.509	1.802	1.126	0.289
	(0.125)	(0.088)	(9)				
5. Electrical goods	0.133	0.848	2.960	0.982	6.043	0.902	0.342
	(0.038)	(0.046)	(10)				
6. Transport equipment[b]	0.308	0.923			2.163		
	(0.182)	(0.061)					
7. Food, drink and tobacco	0.303	0.931	2.415	0.966	2.255	0.295	0.587
	(0.137)	(0.040)	(8)				
8. Textile, leather and shoes	0.066	0.976	1.120	0.999	2.161	0.357	0.550
	(0.097)	(0.035)	(9)				
9. Timber and furniture[b]	0.103	0.932			1.787		
	(0.107)	(0.066)					
10. Paper and printing products	0.227	0.936	3.846	0.871	1.902	1.716	0.190
	(0.080)	(0.036)	(8)				

[a] We trim 5% of observations at the right tail.
[b] We have been unable to compute the second-step GMM estimate.

Table 8: Hicks-neutral technological change.

Industry	$\Delta\omega_H$	$\omega_H{}^a$		Firms' R&D activities			Total technological change	
				ω_H	$\Delta\omega_H$			
		IQR	AC	R&D–No R&D	R&D	No R&D	$\epsilon_{L,-2}\Delta\omega_L + \Delta\omega_H$	$corr(\epsilon_{L,-2}\Delta\omega_L, \Delta\omega_H)^a$
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
1. Metals and metal products	0.044	0.718	0.736	0.004	0.046	0.038	0.065	0.125
2. Non-metallic minerals	0.005	0.551	0.685	0.262	-0.019	0.041	0.036	0.350
3. Chemical products	0.019	0.367	0.872	0.002	0.022	0.011	0.035	-0.054
4. Agric. and ind. machinery	0.041	0.979	0.890	0.446	0.039	0.022	0.073	0.000
5. Electrical goods	0.020	0.697	0.844	0.723	0.009	0.055	0.041	0.024
6. Transport equipment	0.042	0.553	0.596	0.136	0.058	-0.031	0.072	0.350
7. Food, drink and tobacco	0.001	0.761	0.909	-0.146	0.007	0.000	0.007	0.470
8. Textile, leather and shoes	0.012	0.715	0.860	-0.199	-0.003	0.032	0.021	0.304
9. Timber and furniture	0.021[b]	0.772	0.772	-0.098	0.008	0.035	0.023[b]	0.689
10. Paper and printing products	0.002	0.607	0.878	-0.130	0.007	0.006	0.014	0.262
All industries	0.014						0.030	

[a] Without replication and weighting.
[b] We trim values of $\Delta\omega_H$ below -0.25 and above 0.5. This amounts to trimming around one-third of observations.

Table 9: Aggregate productivity growth decomposition.

	Period	Change in aggregate productivity[a,b]	Decomposition				
			Survivors				
			Total	Shift	Covariance	Entrants	Exitors
		(1)	(2)	(3)	(4)	(5)	(6)
$\epsilon_{L,-2}\omega_L$	1992-2006	0.217	0.150	0.130	0.021	0.026	0.041
	1992-1996	0.106	0.096	0.084	0.012	0.017	-0.006
	1997-2001	0.067	0.064	0.066	-0.002	0.006	-0.003
	2002-2006	0.066	0.061	0.074	-0.013	0.000	0.005
ω_H	1992-2006	0.197	0.150	0.151	-0.001	0.030	0.017
	1992-1996	0.098	0.062	0.060	0.002	0.036	0.001
	1997-2001	0.050	0.055	0.056	-0.001	-0.008	0.003
	2002-2006	-0.004	-0.012	-0.005	-0.007	0.001	0.008
$\epsilon_{L,-2}\omega_L + \omega_H$	1992-2006	0.414	0.307	0.284	0.023	0.056	0.051
	1992-1996	0.207	0.164	0.142	0.022	0.048	-0.006
	1997-2001	0.103	0.099	0.100	0.000	-0.002	0.005
	2002-2006	0.059	0.045	0.064	-0.019	0.001	0.014

[a] We trim 1% of observations at each tail of the productivity distribution separately for survivors, entrants, and exitors but pooled across the start and end year.

[b] Changes over the subperiods do not add up because of trimming and because subperiods do not overlap.

Table 10: Capital-augmenting technological change.

Industry	$\Delta\omega_K$	$\epsilon_{K,-2}\,\Delta\omega_K$	GMM				
			β_K (s. e.)	ν (s. e.)	δ_K (s. e.)	χ^2 (df)	p val.
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
1. Metals and metal products	0.056	0.004	0.254	0.903	0.036	2.555	0.923
			(0.129)	(0.055)	(0.061)	(7)	
2. Non-metallic minerals	-0.010	0.007	0.236	0.906	0.010	3.979	0.782
			(0.102)	(0.072)	(0.072)	(7)	
3. Chemical products	-0.018	0.001	0.125	0.942	-0.031	0.598	0.999
			(0.068)	(0.041)	(0.092)	(7)	
4. Agric. and ind. machinery	-0.020	0.000	0.182	0.801	0.031	9.026	0.340
			(0.177)	(0.081)	(0.122)	(8)	
5. Electrical goods	-0.078	0.000	0.129	0.845	-0.004	2.493	0.981
			(0.041)	(0.054)	(0.056)	(9)	
6. Transport equipment[a]	0.008	0.005	0.115	0.981	-0.143		
			(0.088)	(0.050)	(0.138)		
7. Food, drink and tobacco	-0.005	0.002	0.282	0.918	-0.045	2.279	0.943
			(0.286)	(0.058)	(0.204)	(7)	
8. Textile, leather and shoes	-0.085	-0.002	0.080	0.971	0.053	2.714	0.951
			(0.143)	(0.047)	(0.135)	(8)	
9. Timber and furniture[a]	-0.042	0.000	0.088	0.924	-0.021		
			(0.119)	(0.067)	(0.059)		
10. Paper and printing products	0.022	0.007	0.229	0.935	0.005	3.066	0.879
			(0.089)	(0.033)	(0.045)	(7)	
All industries	-0.010	0.003					

[a] We have been unable to compute the second-step GMM estimate.

Table A1: Industry definitions and equivalent classifications.

Industry	Classifications: ESSE	Classifications: National Accounts	Classifications: ISIC (Rev. 4)	Share of value added	Number of subindustries
	(1)	(2)	(3)	(4)	(5)
1. Ferrous and non-ferrous metals and metal products	12+13	DJ	C 24+25	13.2	11
2. Non-metallic minerals	11	DI	C 23	8.2	8
3. Chemical products	9+10	DG-DH	C 20+21+22	13.9	7
4. Agricultural and industrial machinery	14	DK	C 28	7.1	7
5. Electrical goods	15+16	DL	C 26+27	7.5	13
6. Transport equipment	17+18	DM	C 29+30	11.6	7
7. Food, drink and tobacco	1+2+3	DA	C 10+11+12	14.5	10
8. Textile, leather and shoes	4+5	DB-DC	C 13+14+15	7.6	11
9. Timber and furniture	6+19	DD-DN 38	C 16+31	7.0	6
10. Paper and printing products	7+8	DE	C 17+18	8.9	4
All industries				99.5	84

Table A2: Variation in the wage and its determinants.

Industry	Wage: CV	Wage: Var	Wage: Within (%)	Wage: Betw. (%)	Wage regression: Temp. (s. e.)	White (s. e.)	Engin. (s. e.)	Tech. (s. e.)	R^2	R^2_Q	R^2_C
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
1. Metals and metal products	0.425	39.025	9.779	29.246	-0.425	0.127	1.106	0.316	0.651	0.094	0.480
			(25.1)	(74.9)	(0.057)	(0.097)	(0.298)	(0.094)			
2. Non-metallic minerals	0.441	36.252	10.072	26.180	-0.098	0.124	0.896	0.246	0.742	0.020	0.643
			(27.8)	(72.2)	(0.065)	(0.159)	(0.280)	(0.181)			
3. Chemical products	0.440	54.332	9.673	44.659	-0.465	0.461	0.592	0.203	0.755	0.197	0.376
			(17.8)	(82.2)	(0.066)	(0.074)	(0.137)	(0.099)			
4. Agric. and ind. machinery	0.354	30.980	11.472	19.508	-0.273	0.285	0.803	-0.028	0.631	0.082	0.484
			(37.0)	(63.0)	(0.067)	(0.105)	(0.226)	(0.125)			
5. Electrical goods	0.383	31.047	8.461	22.586	-0.374	0.219	1.092	0.312	0.661	0.200	0.356
			(27.3)	(72.7)	(0.058)	(0.073)	(0.264)	(0.087)			
6. Transport equipment	0.393	40.666	12.876	27.790	-0.377	0.220	0.402	0.274	0.709	0.066	0.552
			(31.7)	(68.3)	(0.079)	(0.108)	(0.300)	(0.166)			
7. Food, drink and tobacco	0.502	36.590	5.952	30.638	-0.451	0.115	1.292	0.357	0.753	0.097	0.481
			(16.3)	(83.7)	(0.053)	(0.053)	(0.265)	(0.154)			
8. Textile, leather and shoes	0.449	16.565	3.654	12.911	-0.260	0.646	1.584	0.346	0.683	0.140	0.389
			(22.1)	(77.9)	(0.048)	(0.084)	(0.402)	(0.241)			
9. Timber and furniture	0.392	14.646	3.643	11.003	-0.356	0.173	0.288	0.002	0.697	0.061	0.525
			(24.9)	(75.1)	(0.051)	(0.089)	(0.377)	(0.164)			
10. Paper and printing products	0.464	51.667	10.003	41.664	-0.477	0.188	0.444	0.277	0.702	0.070	0.505
			(19.4)	(80.6)	(0.099)	(0.084)	(0.210)	(0.127)			